

14/12/70

COMPILED BY XALE MK. 4B

```
'DUMP ON' (ED)
'RUN'
'LEADERS'
'OVERLAY PROGRAM' (TEMP)
'BLOCKSIZES'
'OVERLAY' (1) 3:4:10:16:17;
'BEGIN' 'COMMENT' GRAHAM WHITEHEAD FOURIER/HANNING SPECTRUM ANALYSIS;
        'COMMENT' INSERT PARAMETER CARDS
          AUTO:= 1 OR 0
          CROSS:= 1 OR 0
          SU:= TAPE SPEED UP RATIO
              N := NUMBER OF CORRELATION POINTS
              T := TIME DELAY INCREMENT
          AUTONORM := R(INFINITY)
          CROSSNORM := RXY(INFINITY)
              TD:= TIME DELAY FOR RXY(T);
          'REAL' A,D,NORM,RX0,C,PIE,AUTONORM,CROSSNORM,T,TD,P,
DCK      1
          B,Q;
          'INTEGER' K,N,J,F,H,Z,E,SU,S1,S2,G,Y;
          'BOOLEAN' AUTO,CROSS;
AUTO:= 'TRUE';
CROSS:= 'TRUE';
        N:=256;
        SU:=4;
        Z:=1;
        T:=0.02;
        PIE := 3.14159265358979323846;
        P:=0;
'BEGIN' 'REAL' 'ARRAY' RX[1:N],RXY[1:N],MCOS[1:N],MSIN[1:N],
DCK      2
        NSIN[1:N], NCOS[1:N],M[1:N];
'PROCEDURE' PART 1;
DCK      3
'BEGIN'
        L1: 'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
'BEGIN'
        MCOS[F]:=0.0;
        MSIN[F]:=0.0;
'END';
        SELECT INPUT(3);
        TD:=READ;
        Q:=PEAD;
        'COMMENT' PRESENT AUTO DATA FIRST THEN CROSS;
        'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO' RX[F]:=0.0;
        A:= READ;
        'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
'BEGIN'
        A:= READ;
        D:= READ;
        RX[F]:= RX[F] + D;
'END';
        A:=READ;
        A:=READ;
        A:=READ;
        'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO' RXY[F]:=0.0;
        'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
```



```

'BEGIN'
      A:= READ;
      D:= READ;
      RXY[F]:= RXY[F] + D ;
'END';
      B:=0.0;
      'FOR' F:=1 'STEP' 1 'UNTIL' 100 'DO'
'BEGIN'
      'IF' RX[F]> B 'THEN'
'BEGIN'
      B:=RX[F];
      Y:=F;
'END'
'END';
      RX0:=RX[Y];
      A:=0.0;
      'FOR' F:=1 'STEP' 1 'UNTIL' 100 'DO'
'BEGIN'
      'IF' ABS(RXY[F])> A 'THEN'
'BEGIN'
      A:=ABS(RXY[F]);
      Z:= F;
'END'
'END';
      CROSSNORM:=0.0;
      AUTONORM:=0.0;
      'IF' Z>Y 'THEN'
'BEGIN'
      'FOR' F:=(Y+P)'STEP'1'UNTIL'N-(Z-Y)'DO' AUTONORM:=
      AUTONORM+RX[F];
      'FOR' F:=(Z+P)'STEP'1'UNTIL'N'DO' CROSSNORM
      := CROSSNORM + RXY[F];
      AUTONORM:=AUTONORM/(N-(Z+P-1));
      CROSSNORM:=CROSSNORM/(N-(Z+P-1));
      G:=(N-Z)+1;
'END'
      'ELSE'
'BEGIN'
      'FOR' F:=(Y+P)'STEP'1'UNTIL'N'DO' AUTONORM :=
      AUTONORM + RX[F];
      'FOR' F:=(Z+P)'STEP'1'UNTIL'N-(Y-Z)'DO' CROSSNORM
      :=CROSSNORM + RXY[F];
      AUTONORM:=AUTONORM/(N-(Y+P-1));
      CROSSNORM:=CROSSNORM/(N-(Y+P-1));
      G:=(N-Y)+1;
'END';
      'IF' (A- CROSSNORM)>(B- AUTONORM) 'THEN'
'BEGIN'
      'FOR' F:=1'STEP' 1'UNTIL' N 'DO'
'BEGIN'
      RX[F]:=(RX[F]- AUTONORM)/(A- CROSSNORM);
      RXY[F]:=(RXY[F] - CROSSNORM)/(A- CROSSNORM);
'END'
'END'
      'ELSE'
'BEGIN'
      'FOR' F:=1'STEP'1'UNTIL'N 'DO'
'BEGIN'
      RX[F]:=(RX[F]- AUTONORM)/( B- AUTONORM);
      RXY[F]:=(RXY[F]- CROSSNORM)/( B- AUTONORM);

```



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'END'
'END'
      'FOR' F:=1 'STEP' 1 'UNTIL' G 'DO'
      M[F]:=FRX[(F+Y)-1];
      E:=ENTIER(G*Q);
      H:= 0;
'END' OF PART 1 PROCEDURE;
'PROCEDURE' PART 2 (LAB); 'LABEL' LAB;
CK      4
'BEGIN'
  RESTART:      H:=H+1;
  'COMMENT' THE NEXT SEGMENT CALCULATES THE FOURIER COEFFICIENTS
              AT SPECIAL FREQUENCIES (2J+1)*2PIE/NT THE RESULT IS
              DEPOSITED IN ARRAYS MSIN AND MCOS;
'BEGIN'
      J:= -1;
  RECYCLE:      J:= J+1;
                K:=0.0;
  LOOP :        K:=K+1;
                A:=(J/2)*(K-1)/(N-1);
                A:=A-ENTIER(A);
                'IF' A>0.0 'AND' A <0.25 'THEN'
'BEGIN'
                MSIN[J+1]:=MSIN[J+1]+M[K]*SIN(2*PIE*A);
                MCOS[J+1]:=MCOS[J+1]+M[K]*COS(2*PIE*A);
'END'
                'ELSE' 'IF' A>0.25 'AND' A <0.5 'THEN'
'BEGIN'
                MSIN[J+1]:= MSIN[J+1]+ M[K]*SIN(PIE-2*PIE*A);
                MCOS[J+1]:=MCOS[J+1]+M[K]*(-1)*COS(PIE-2*PIE*A);
'END'
                'ELSE' 'IF' A>0.5 'AND' A<0.75 'THEN'
'BEGIN'
                MSIN[J+1]:= MSIN[J+1] +M[K]*SIN((-1)*(2*PIE*A -PIE));
                MCOS[J+1]:= MCOS[J+1] +M[K]*COS((-1)*(2*PIE*A -PIE));
'END'
                'ELSE' 'IF' A>0.75 'AND' A < 1.0 'THEN'
'BEGIN'
                MSIN[J+1]:= MSIN[J+1]+M[K]*SIN((-1)*(2*PIE-2*PIE*A));
                MCOS[J+1]:= MCOS[J+1]+M[K]*COS(      (2*PIE-2*PIE*A));
'END'
                'ELSE' 'IF' A=0.0 'OR' A=1.0 'THEN'
                MCOS[J+1]:= MCOS[J+1] +M[K]
'ELSE' 'IF' A=0.5 'THEN' MCOS[J+1]:=MCOS[J+1]+M[K]*(-1)
'ELSE' 'IF' A=0.25 'THEN' MSIN[J+1]:=MSIN[J+1]+M[K]
'ELSE' 'IF' A=0.75 'THEN' MSIN[J+1]:= MSIN[J+1]+M[K]*(-1);
'END';
                'IF' K = G 'THEN' 'GOTO' L6 'ELSE' 'GOTO' LOOP;
      L6:      'IF' J=( E -1) 'THEN' 'GOTO' L7 'ELSE' 'GOTO' RECYCLE ;
      L7:      MSIN[1]:=0.5*MSIN[1]+0.5*MSIN[2];
                MCOS[1]:=0.5*MCOS[1]+0.5*MCOS[2];
                'FOR' J:=2 'STEP' 1 'UNTIL' (E-1) 'DO'
'BEGIN'
      MSIN[J]:=0.25*MSIN[J-1]+0.5*MSIN[J]+0.25*MSIN[J+1];
      MCOS[J]:=0.25*MCOS[J-1]+0.5*MCOS[J]+0.25*MCOS[J+1];
'END';
      MSIN[E]:=0.5*MSIN[E-1]+0.5*MSIN[E];
      MCOS[E]:=0.5*MCOS[E-1]+0.5*MCOS[E];
                'IF' H = 1 'AND' AUTO 'THEN' 'GOTO' L8 'ELSE' 'IF' H=1
                'AND' 'NOT' AUTO 'THEN' 'GOTO' L9 'ELSE' 'GOTO' L10;

```



```

L8:      'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
'BEGIN'
          NSINE[F] := MSINE[F];
          NCOS[F] := MCOS[F];
'END';
      'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
'BEGIN'
          MSINE[F]:=0.0;
          MCOS[F]:=0.0;
'END';
          'IF' AUTO 'AND' CROSS 'THEN' 'GOTO' L11 'ELSE' 'GOTO'
          L10;
L11:      'FOR' F:=1 'STEP' 1 'UNTIL' G 'DO'
          MLF] := RXY[(F+Z)-1];
          'GOTO' RESTART;
L9 :      'GOTO' L13;
L10: 'COMMENT' THE FOLLOWING PROCEEDURE PLOTS THE AUTOCORRELOGRA
          AND THE POWER SPECTRAL DENSITY;
'END' OF PART 2;
'BEGIN'
          'REAL' MAX,POWER;
CK      5
          'INTEGER' LITA;
          'REAL' 'ARRAY' AB[1:N],BC[1:N] ;
          'INTEGER' 'ARRAY' CD[1:17],DE[1:3],EF[1:22],FG[1:13],
          GH[1:13],HI[1:24],JK[1:28],KL,MN,TU,VW[1:4],LM[1:5],
          NO,OP,PQ,QR[1:7],RS[1:20],ST[1:25],UV[1:9];
'BEGIN'
          'PROCEDURE' OPENPLOT;'EXTERNAL';
CK      6
CK      7
          'PROCEDURE' CLOSEPLOT;'EXTERNAL';
CK      8
          'PROCEDURE' HG PLOT (X,Y,IC,L);'VALUE'X,Y,IC,L;'INTEGER'IC,L;
CK      9
          'REAL'X,Y;'EXTERNAL';
          'PROCEDURE' STRARR(A,N,S);'ARRAY'A;'INTEGER' N;
CK     10
          'STRING' S; 'EXTERNAL';
          'PROCEDURE' HGPAXISV (X,Y,BCD,NC,S,THETA,XMIN,DX,GAP,NH);
CK     11
          'VALUE'X,Y,NC,S,THETA,XMIN,DX,GAP,NH;'INTEGER'NC,NH;'ARRAY'
          BCD;'REAL'X,Y,S,THETA,XMIN,DX,GAP;'EXTERNAL';
          'PROCEDURE' HG PLINE (X,Y,N,K);'VALUE'N,K;'ARRAY'X,Y;
CK     12
          'INTEGER' N,K;'EXTERNAL';
          'PROCEDURE' HGPNUMBER (X,Y,HT,FL,THETA,I,IP,IQ);'VALUE'X,Y,
CK     13
          HT,FL,THETA,I,IP,IQ;'INTEGER'I,IP,IQ;'REAL'X,Y,HT,FL,THETA;
          'EXTERNAL';
          'PROCEDURE' HG PSYMBL (X,Y,HT,BCD,THETA,N);'VALUE'X,Y,HT,
CK     14
          THETA,N;'INTEGER'N;'ARRAY'BCD;'REAL'X,Y,HT,THETA;'EXTERNAL';
'PROCEDURE' PART 3;
CK     15
'BEGIN'
          OPENPLOT ;
          HG PLOT (-3.0,10.5,0,4);
          MAX:=0.0;
          'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

```



```
'BEGIN'
  'IF' ABS( NCOS[J ]) > MAX 'THEN' MAX:=ABS( NCOS[J]) 'ELSE' MAX:=MAX;
OUTPUT( NCOS[J] );
'END';
```

```
CROSSNORM:=MAX;
```

```
NEWLINE(1);
OUTPUT( MAX );
```

```
'COMMENT' THE NEXT SEGMENT COMPUTES THE X AND Y COORDINATES TO
PLOT A HISTOGRAM;
```

```
A:=(2*PIE)/((G+1)*T*SU);
```

```
D:=(ENTIER(((PIE/T)/SU)/8)+1)*8*Q;
```

```
'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'
```

```
'BEGIN'
```

```
AB[1]:= (J-1)*(A/2)*4.5/D;
```

```
AB[2]:= AB[1];
```

```
AB[3]:= AB[1];
```

```
BC[1]:= 0.0;
```

```
BC[2] := ABS( NCOS[J]) * 2.5/MAX;
```

```
BC[3]:= 0.0;
```

```
HGPLINE (AB, BC, 3, +1);
```

```
OUTPUT( NSIN[J] );
```

```
'END';
```

```
'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'
```

```
'BEGIN'
```

```
AB[J]:= (J-1)*(A/2)*4.5/D;
```

```
BC[J]:= ABS( NCOS[J]) * 2.5/MAX;
```

```
'END';
```

```
HGPLINE( AB, BC, E, +1 );
```

```
STRARR( CD, LITA, '( 'FREQUENCY%RAD/SEC' ) );
```

```
HGPAXISV ( 0.0, 0.0, 0.0, CD, -17, 4.5, 0.0, 0.0, 0.0, D/10, 0.45, +2 );
```

```
STRARR( DE, LITA, '( 'XXX' ) );
```

```
HGPAXISV ( 0.0, 0.0, 0.0, DE, 0.3, 5.90, 0.0, 0.0, 0.5, 1.25, +2 );
```

```
STRARR( EF, LITA, '( 'POWER%SPECTRAL%DENSITY' ) );
```

```
HGPSYMBL ( =1.0, 0.5, 0.125, EF, 90, 0, 22 );
```

```
C:=(4.5*2*PIE)/D;
```

```
STRARR( EF, LITA, '( 'CYCLES/SECOND' ) );
```

```
HGPAXISV ( 0.0, -0.5, EF, -13, 4.5, 0.0, 0.0, 0.0, 1.0, C, -1 );
```

```
STRARR( FG, LITA, '( 'CM, CM/RAD/SEC' ) );
```

```
HGPSYMBL ( =0.625, 1.0, 0.1, FG, 90, 0, 13 );
```

```
STRARR( GH, LITA, '( 'G. D. WHITEHEAD' ) );
```

```
HGPSYMBL ( 3.5, 8.875, 0.1, GH, 0, 0, 13 );
```

```
STRARR( HJ, LITA, '( 'FOURIER/HANNING%TRANSFORM' ) );
```

```
HGPAXISV ( 0.0, 6.0, DE, -0.4, 5.0, 0.0, 0.0, (T*(N-1)*SU)/5, 0.9, +1 );
```

```
HGPSYMBL ( -1.25, 8.875, 0.2, HJ, 0, 0, 25 );
```

```
STRARR( JK, LITA, '( 'AUTOCORRELATION%COEFFICIENT' ) );
```

```
HGPSYMBL ( =1.0, 4.75, 0.125, JK, 90, 0, 27 );
```

```
HGPAXISV ( 0.0, 4.25, DE, 0.3, 5.90, 0.0, =1.0, 0.5, 0.875, +1 );
```

```
HGPAXISV ( 0.0, 6.0, DE, -0.4, 5.0, 0.0, 0.0, (T*(N-1)*SU)/5, 0.9, +1 );
```

```
STRARR( KL, LITA, '( 'TIME' ) );
```

```
HGPSYMBL ( 4.5, 6.125, 0.1, KL, 0, 0, 4 );
```

```
STRARR( LM, LITA, '( 'DELAY' ) );
```

```
HGPSYMBL ( 4.5, 5.75, 0.1, LM, 0, 0, 5 );
```

```
STRARR( MN, LITA, '( 'SECS' ) );
```

```
HGPSYMBL ( 4.5, 5.5, 0.1, MN, 0, 0, 4 );
```

```
HGPLOT ( 0, 0, 0, 0, 3, 0 );
```

```
'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
```

```
M[F]:= (F/N*4.5) - (4.5/N);
```

```
'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
```

```
RX[F]:= 1.75 * RX[F];
```

```
'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
```



```

RXY[FJ]=1.75*RXY[F];
HG PLOT (0.0,-6.0,0.4);
HG PLINE (M,RX,N,+1);
STRARR(NO,LITA,('R(O)%%='));
HG PSYMBL (2.875,1.65,0.1,NO,0.0,7);
STRARR(OP,LITA,('R(INF)='));
HG PSYMBL (2.875,1.5,0.1,OP,0.0,7);
HG PNUMBER (3.375,1.65,0.1,RX0,0.0,1.0,2,6);
HG PNUMBER (3.375,1.50,0.1,AUTONORM,0.0,1.0,2,6);
STRARR(PQ,LITA,('POWER='));
POWER:=0.0;
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
POWER:=POWER+NCOS[F];
HG PSYMBL (2.875,-2.625,0.1,PQ,0.0,6);
HG PNUMBER (3.125,-2.625,0.1,POWER,0.0,1.0,2,6);
STRARR(QR,LITA,('%MAX='));
HG PSYMBL (2.875,-2.475,0.1,QR,0.0,6);
HG PNUMBER (3.125,-2.475,0.1,MAX,0.0,1.0,2,6);
AB[1] := 5.25;
AB[2] := 5.25;
AB[3] := -1.75;
AB[4] := -1.75;
AB[5] := 5.25;
BC[1] := -7.25;
BC[2] := +3.25;
BC[3] := +3.25;
BC[4] := -7.25;
BC[5] := -7.25;
HG PLINE (AB,BC,5,+1);
HG PLOT (0.0,0.0,3,0);
L11: HG PLOT (-10.75,0.0,0.4);
HG PLINE (M,RXY,N,+1);
HG PAXISV (0.0,-1.75,DE,0,3,5,90.0,-1,0.5,0.875,+2);
HG PAXISV (0.0,0.0,DE,-0,4,5,0.0,0.0,(T*(N+1)*SU)/5,0.9,+1);
STRARR(JK,LITA,('CROSSCORRELATION%COEFFICIENT'));
HG PSYMBL (-1.0,-1.375,0.125,JK,90,0,28);
HG PSYMBL (4.625,0.125,0.1,KL,0.0,4);
HG PSYMBL (4.625,0.0,0.1,LM,0.0,5);
HG PSYMBL (4.625,-0.125,0.1,MN,0.0,4);
MAX:=0.0;
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
'BEGIN;
'IF' ABS(MSIN[F]) > MAX 'THEN' MAX :=ABS(MSIN[F]) 'ELSE' MAX:=MAX;
'END';
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
'BEGIN;
'IF' ABS(MCOS[F]) > MAX 'THEN' MAX :=ABS(MCOS[F]) 'ELSE' MAX
:=MAX;
'END';
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
'BEGIN;
RX[F] := (MCOS[F]/MAX)*2.5;
RXY[FJ] := (MSIN[F]/MAX)*2.5;
'END';
HG PLOT (0,0,0,0,3,0);
HG PLOT (-1,75,4.625,0,4);
HG PLINE (RX,RXY,E,+1);
HG PAXISV (-2.5,0.0,DE,0,5,0,0.0,-1,0,0,2,0.5,1);
HG PAXISV (0.0,-2.5,DE,0,5,0,90.0,-1,0,0,2,0.5,1);

```



```

STRARR(UV,LITA,('IMAGINARY'));
HGPSYMBL (2.75,0.0,0.125,YU,0.0,4);
HGPSYMBL (-0.5,2.7,0.125,UV,0.0,9);
STRARR(RS,LITA,('FOURIER/HANNING%GROSS'));
STRARR(ST,LITA,('SPECTRAL%DENSITY%ANALYSIS'));
HG PLOT (0,0,0,0,3,0);
HG PLOT (0,0,-1,875,0,4);
HGPSYMBL (1.5,5.5,0.1, GH,0.0,13);
HGPSYMBL (-2.875,5.5,0.125,RS,0.0,21);
HGPSYMBL (-2.875,5.25,0.125,ST,0.0,25);
STRARR(VW,LITA,('MAX=1'));
HGPSYMBL (-2.875,5.0,0.1,VW,0.0,4);
HGPNUMBER (-2.275,5.0,0.1,MAX,0.0,1,0,6);
AB[1] := -3.5;
AB[2] := -3.5;
AB[3] := 3.5;
AB[4] := 3.5;
AB[5] := -3.5;
BC[1] := 6.0;
BC[2] := -6.5;
BC[3] := -6.5;
BC[4] := 6.0;
BC[5] := 6.0;
HG PLINE (AB,BC,5,+1);

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'END' OF PART 3 ;

'PROCEDURE' PART 4 ;

CK 16

'BEGIN'

'COMMENT' THE NEXT SEGMENT PLOTS THE CO AND QUAD SPECTRUM;

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HG PLOT (0,0,0,0,3,0);
HG PLOT (-10.0,0.0,0,4);
HGPSYMBL (-2.875,5.5,0.125,RS,0.0,21);
HGPSYMBL (-2.875,5.25,0.125,ST,0.0,25);
HGPSYMBL (1.5,5.5,0.1, GH,0.0,13);
HG PLINE (AB,BC,5,+1);

```

```

HG PLOT (0,0,0,0,3,0);
HG PLOT (1.75,2.625,0,4);

```

'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

'BEGIN'

```

AB[1]:=(J-1)*(A/2)*4.5/D;
AB[2]:=AB[1];
AB[3]:=AB[1];
BC[1]:=0,0;
BC[2]:=MSIN[J]*2.5/MAX;
BC[3]:=0,0;
HG PLINE (AB,BC,3,+1);

```

'END';

'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

'BEGIN'

```

AB[J]:=(J-1)*(A/2)*4.5/D;
BC[J]:=MSIN[J]*2.5/MAX;

```

'END';

```

HG PLINE(AB,BC,E,+1);

```

AB[1]:=0,0;

AB[2]:=4,5;

BC[1]:=0,0;

BC[2]:=0,0;

HG PLINE(AB,BC,2,+1);

HG PAXISV (0,0,0,0,CD,-17,4.5,0.0,0.0,D/10,0.45,-2);


```

HGPAXISV (0,0,-1.25,DE,0,3.75,90,0,-0.5,0.5,1.25,*2);
STRARR(JK,LITA,('QUADXSPECTRUM')));
HGPSYMBL (-1.0,-0.25,0.125,JK,90,0,13);
HGPSYMBL (-0.625,-0.25,0.1,FG,90,0,13);
HG PLOT (0,0,0,0,3,0);
HG PLOT (0,0,-4.375,0,4);

```

```

'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

```

'BEGIN'

```

AB[1]:=(J-1)*(A/2)*4.5/D;
AB[2]:=AB[1];
AB[3]:=AB[1];
BC[1]:=0,0;
BC[2]:=MCOS[J]*2.5/MAX;
BC[3]:=0,0;
HG PLINE (AB,BC,3,+1);

```

'END'

```

'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

```

'BEGIN'

```

AB[J]:=(J-1)*(A/2)*4.5/D;
BC[J]:=MCOS[J]*2.5/MAX;

```

'END'

```

HG PLINE(AB,BC,E,+1);
AB[1]:=0,0;
AB[2]:=4,5;
BC[1]:=0,0;
BC[2]:=0,0;
HG PLINE(AB,BC,2,+1);

```

```

HGPAXISV (0,0,0,0,CD,-17,4.5,0,0,0,0,D/10,0.45,*2);
HGPAXISV (0,0,-1.25,EF,13,4.5,0,0,0,0,1,0,C,-1);
HGPAXISV (0,0,-1.25,DE,0,3.75,90,0,-0.5,0.5,1.25,*2);
STRARR(JK,LITA,('COXSPECTRUM')));
HGPSYMBL (-1,0,0,0,0.125,JK,90,0,11);
HGPSYMBL (-0.625,-0.25,0.1,FG,90,0,13);

```

'END' OF PART 4;

'PROCEDURE' PART 5;

OK 17

'BEGIN'

'COMMENT' THE NEXT SEGMENT COMPUTES AND PLOTS THE GAIN & PHASE;

```

'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
RX[F]:=20,0*0.43429448*(LN(SQRT((MSIN[F]*MSIN[F]
+(MCOS[F]*MCOS[F]))/
SQRT((NSIN[F]*NSIN[F])+(NCOS[F]*NCOS[F]))));
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
RX[F]:=RX[F]*1.675/20;
HG PLOT (0,0,0,0,3,0);
HG PLOT (10,0,14.25,0,4);

```

```

'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

```

'BEGIN'

```

AB[1]:=(J-1)*(A/2)*4.5/D;
AB[2]:=AB[1];
AB[3]:=AB[1];
BC[1]:=0,0;
BC[2]:=RX[J];
BC[3]:=0,0;
HG PLINE (AB,BC,3,+1);

```

'END'

```

'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

```

```

AB[J]:=(J-1)*(A/2)*4.5/D;
HG PLINE(AB,BC,E,+1);

```



```

HGPAXISV (0,0,0,0,CD,-17,4.5,0,0,0,0,0,0,0,10,0.45,-2);
HGPAXISV (0,0,-1.875,EF,13,4.5,0,0,0,0,0,1,0,C,-1);
HGPAXISV (0,0,-1.875,DE,0,3.75,90,0,-20,10,0.9375,-2);
STRARR(JK,LITA,('GAIN-DECIBELS'));
HGPSYMBL (-1,0,-0.875,0.125,JK,90,0,13);
HG PLOT (0,0,0,0,3,0);
HG PLOT (0,0,4.375,0,4);
HGNUMBER (2.25,1,1,0.1,TD,0,0,0,2,3);
STRARR(JK,LITA,('MXSECS. '));
HGPSYMBL (3,0,1,1,0.1,JK,0,0,7);
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
RX[F]:=-(180*TD+F)/(T*(G-1)*SU);
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'

```

'BEGIN'

```

S1:=SIGN(MCOS[F]);
S2:=SIGN(MSIN[F]);
T:=ABS(ARCTAN(MSIN[F]/MCOS[F])*57.29578);
'IF' S1=1 'AND' S2=-1 'AND' F<17 'THEN' RX[F]:=RX[F]+T
'ELSE' 'IF' S1=0 'AND' S2=0 'THEN' RX[F]:=RX[F]+0.0
'ELSE' 'IF' S1=1 'AND' S2=1 'THEN' RX[F]:=RX[F]-T
'ELSE' 'IF' S1=0 'AND' S2=1 'THEN' RX[F]:=RX[F]-90
'ELSE' 'IF' S1=0 'AND' S2=-1 'THEN' RX[F]:=RX[F]+270
'ELSE' 'IF' S1=1 'AND' S2=0 'THEN' RX[F]:=RX[F]+0.0
'ELSE' 'IF' S1=-1 'AND' S2=0 'THEN' RX[F]:=RX[F]-180
'ELSE' 'IF' S1=-1 'AND' S2=1 'THEN' RX[F]:=RX[F]-(180-T)
'ELSE' 'IF' S1=1 'AND' S2=-1 'THEN' RX[F]:=RX[F]-(360-T)
'ELSE' 'IF' S1=-1 'AND' S2=-1 'THEN' RX[F]:=RX[F]-(180+T);

```

'END'

```
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
```

'BEGIN'

```

S1:=SIGN(NCOS[F]);
S2:=SIGN(NSIN[F]);
T:=ABS(ARCTAN(NSIN[F]/NCOS[F])*57.29578);
'IF' S1=1 'AND' S2=-1 'AND' F<17 'THEN' RX[F]:=RX[F]-T
'ELSE' 'IF' S1=0 'AND' S2=0 'THEN' RX[F]:=RX[F]+0.0
'ELSE' 'IF' S1=1 'AND' S2=1 'THEN' RX[F]:=RX[F]+T
'ELSE' 'IF' S1=0 'AND' S2=1 'THEN' RX[F]:=RX[F]+90
'ELSE' 'IF' S1=0 'AND' S2=-1 'THEN' RX[F]:=RX[F]+270
'ELSE' 'IF' S1=1 'AND' S2=0 'THEN' RX[F]:=RX[F]+0.0
'ELSE' 'IF' S1=-1 'AND' S2=0 'THEN' RX[F]:=RX[F]+180
'ELSE' 'IF' S1=-1 'AND' S2=1 'THEN' RX[F]:=RX[F]+(180-T)
'ELSE' 'IF' S1=1 'AND' S2=-1 'THEN' RX[F]:=RX[F]+(360-T)
'ELSE' 'IF' S1=-1 'AND' S2=-1 'THEN' RX[F]:=RX[F]+(180+T);
'IF' RX[F]<-360 'THEN' RX[F]:=-360;

```

'END'

```
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO' RX[F]:=RX[F]*1.875/360;
```

```
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
```

```
RXY[F]:=F*A*2.25/D;
```

```
HG PLINE (RXY,RX,E,+1);
```

```
HGPAXISV (0,0,0,0,CD,-17,4.5,0,0,0,0,0,0,0,10,0.45,-2);
```

```
HGPAXISV (0,0,0,0,DE,-0,4.5,0,0,0,0,0,0,0,1,0,C,-1);
```

```
HGPAXISV (0,0,-1.875,EF,13,4.5,0,0,0,0,0,1,0,C,-1);
```

```
HGPAXISV (0,0,-1.875,DE,0,3.75,90,0,-360,180,0.9375,-2);
```

```
STRARR(JK,LITA,('PHASE ANGLE-DEGREES'));
HGPSYMBL (-1,0,-1,0,0.125,JK,90,0,20);
```

```
STRARR(JK,LITA,('AGGREGATE TIME DELAY'));
HGPSYMBL (2,25,1,25,0.1,JK,0,0,20);
```

```
HG PLOT (0,0,0,0,3,0);
```

```
HG PLOT (-1.75,-1.75,0,4);
```



```

HOPSYMBL (-2.875,5.25,0.125,ST,0.0,25);
HOPSYMBL (1,5,5,5,0.1,GH,0.0,13);

```

```

AB[1] := -3.5;
AB[2] := -3.5;
AB[3] := 3.5;
AB[4] := 3.5;
AB[5] := -3.5;

```

```

BC[1] := 6.0;
BC[2] := -4.5;
BC[3] := -4.5;
BC[4] := 6.0;
BC[5] := 6.0;
HGPLINE (AB,BC,5,41);

```

```

L12: CLOSEPLOT;

```

```

SELECT OUTPUT(4);

```

```

RUNOUT;

```

```

RUNOUT;

```

```

NEWLINE(1);

```

```

'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'

```

```

'BEGIN'

```

```

    PRINT(NCOS[F],2,5);

```

```

NEWLINE(1);

```

```

'END';

```

```

RUNOUT;

```

```

RUNOUT;

```

```

NEWLINE(1);

```

```

'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'

```

```

'BEGIN'

```

```

    PRINT(NSIN[F],2,5);

```

```

NEWLINE(1);

```

```

'END';

```

```

RUNOUT;

```

```

RUNOUT;

```

```

NEWLINE(1);

```

```

'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'

```

```

'BEGIN'

```

```

    PRINT(MCOS[F],2,5);

```

```

NEWLINE(1);

```

```

'END';

```

```

RUNOUT;

```

```

RUNOUT;

```

```

NEWLINE(1);

```

```

'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'

```

```

'BEGIN'

```

```

    PRINT(MSIN[F],2,5);

```

```

NEWLINE(1);

```

```

'END';

```

```

RUNOUT;

```

```

RUNOUT;

```

```

'END' OF PART 5;

```

```

PART 1;

```

```

PART 2 (L13);

```

```

PART 3;

```

```

PART 4;

```

```

PART 5;

```

```

'END';

```

```

'END';

```

```

L13: 'END';

```



```

BLOCK      8 -      0 WORDS
BLOCK      9 -      0 WORDS
BLOCK     10 -      0 WORDS
BLOCK     11 -      0 WORDS
BLOCK     12 -      0 WORDS
BLOCK     13 -      0 WORDS
BLOCK     14 -      0 WORDS
BLOCK      6 -     41 WORDS
BLOCK      5 -    176 WORDS
BLOCK      2 -     90 WORDS
BLOCK      1 -     22 WORDS
BLOCK      3 -    835 WORDS
BLOCK      4 -   1094 WORDS
BLOCK     15 -   2505 WORDS
BLOCK     16 -   1065 WORDS
BLOCK     17 -   2663 WORDS
NO OF BUCKETS USED      209
COMPILED      #TEMP      EC

```

```

10/46/59  0#INFP  #UAGB
10/47/03  0#INFP  DLTG      OK
10/47/03  0#INFP  CLGD      0
10/47/06  0#XALE  #INFP
10/48/37  0#XALE  DISP  COMPILED #TEMP
10/48/38  0#XALE  DLTG  FI #XPCK
10/48/38  0#XALE  CLGD   10
10/48/41  0#XPCK  #XALE
10/50/10  0#XPCK  DLTG  FI #TEMP
10/50/10  0#XPCK  CLGD   23
10/50/15  0#TEMP  #XPCK
10/54/12  0#TEMP  HALT   77
10/54/14  0#TEMP  MT32   2 *0005555 SCRATCH TAPE  0 *00000000
10/54/21  0#TEMP  MT32   2 *0005555 PLOTTERFILE  0 *00000000
10/58/04  0#TEMP  CLOSED MT32
10/58/04  0#TEMP  HALT  END OF GRAPH
10/58/08  0#TEMP  HALT  AH
10/58/08  0#TEMP  CLGD  216

```

```

JOB HANNING,ASSTF215,WHITEHEAD,STREAM 3      RUN AT ASTON UNIV.
BULLETIN HANNING ,3,300
VOLUME 5000
ALGOLUB
***

```

```

JOB COST NP      178      DATE      14/12/70
JOB NAME      HANNING      START TIME  10/46/49
USER NAME     WHITEHEAD    END TIME   10/58/13
PERIPHERALS USED; 32

```

```

TOTAL MILL TIME      271
INPUT RECORDS      1622
OUTPUT RECORDS      1034

```


14/12/70

COMPILED BY XALE MK. 4B

'DUMP ON' (ED)
'RUN'
'LEADERS'

'LIBRARY' (ED, SUBGROUP GRAH)

'OVERLAY PROGRAM' (TEMP)
'BLOCKSIZES'

'OVERLAY' (1) 3:4:15:16:17;

'BEGIN' 'COMMENT' GRAHAM WHITEHEAD FOURIER/MILNER SPECTRUM ANALYSIS;
'COMMENT' INSERT PARAMETER CARDS

AUTO:= 1 OR 0

CROSS:= 1 OR 0

SU:= TAPE SPEED UP RATIO

N := NUMBER OF CORRELATION POINTS

T := TIME DELAY INCREMENT

AUTONORM := R(INFINITY)

CROSSNORM := RXY(INFINITY)

TD:= TIME DELAY FOR RXY(T);

'REAL' A,D,NORM,RX0,C,PIE,AUTONORM,CROSSNORM,T,TD,P;

CK 1

'INTEGER' K,N,J,F,H,Z,E,SU,S1,S2,G;

'BOOLEAN' AUTO,CROSS;

AUTO:='TRUE';

CROSS:='TRUE';

N:=250;

SU:=4;

T:=0.02;

Z:=1;

P:=42;

PIE := 3.14159265358979323846;

'BEGIN' 'REAL' 'ARRAY' RX[1:N],RXY[1:N],MCOS[1:N],MSINE[1:N],

CK 2

NSIN[1:N], NCOS[1:N],M[1:N];

'PROCEDURE' PART 1;

CK 3

'BEGIN'

L1: 'FOR' F:=1'STEP'1'UNTIL'N'DO'

'BEGIN'

MCOS[F]:=0.0;

MSINE[F]:=0.0;

'END';

SELECT INPUT(3);

TD:=READ;

'COMMENT' PRESENT AUTO DATA FIRST THEN CROSS;

'FOR' F:=1'STEP'1'UNTIL'N'DO' RX[F]:=0.0;

A:= READ;

'FOR' F:=1'STEP'1'UNTIL' N 'DO'

'BEGIN'

A:= READ;

D:= READ;

RX[F]:= RX[F] + D;

'END';

A:=READ;

A:=READ;


```

A:=READ;
DIM READ;
RXY[F]:=RXY[F]+D;
'END';

A:=0.0;
'FOR' F:=P 'STEP' 1 'UNTIL' N 'DO' A:=A+RXY[F];
AUTONORM:=A/(N-(P-1));
A:=0.0;
'FOR' F:=P 'STEP' 1 'UNTIL' N 'DO' A:=A+RXY[F];
CROSSNORM:=A/(N-(P-1));
'COMMENT' NORM IS THE VALUE R(0)-R(INF) THAT IS USED TO NORMAL
RXY AND RX0 IS THE VALUE OF R(0);
NORM := RX[1]-AUTONORM;
RAD := RX[1];
'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'

'BEGIN'
R[X[F] :=(RX[F]-AUTONORM)/(RX0-AUTONORM);
RAY[F] :=(RXY[F]-CROSSNORM)/NORM;
'END';

A:=0.0;
'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
'BEGIN'
'IF' ABS(RXY[F])>A 'THEN'
'BEGIN'
A:=ABS(RXY[F]);
Z:=F;
'END'
'END';

E:=ENTIER(((N-Z)+1)/4);
G:=(N*Z)+1;
'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
ML[F] := RX[F];
'COMMENT' THE NEXT SEGMENT CALCULATES THE MILNER MODULATORS
RSINX/X THE CORRELATION NUMBERS ARE HELD IN
ARRAY M[ ];
HI:=0;
'END' OF PART 1 PROCEDURE;
'PROCEDURE' PART 2 (LAB): 'LABEL' LAB;
CK 4
'BEGIN'
RESTART: HI:=HI+1;
'FOR' K:=1 'STEP' 1 'UNTIL' (G+1)'DO'
'BEGIN'
C:=K/(G-1);
'IF' C>0.0 'AND' C<0.25 'THEN'
M[K+1]:=M[K+1]*SIN(2*PIE*(K/(G-1)))/(K*T) 'ELSE' 'IF'
C>0.25 'AND' C<0.5 'THEN'
M[K+1]:=M[K+1]*SIN(PIE-2*PIE*K/(G-1))/(K*T) 'ELSE' 'IF'
C>0.5 'AND' C<0.75 'THEN'
M[K+1]:=M[K+1]*SIN((-1)*(2*PIE*(K/(G-1))-PIE))/(K*T) 'ELSE'
'IF' C>0.75 'AND' C<1.0 'THEN'
M[K+1]:=M[K+1]*SIN((-1)*(2*PIE-2*PIE*(K/(G-1)))/(K*T)
'ELSE' 'IF' C=0.25 'THEN' M[K+1]:=M[K+1]*(1/(K*T))
'ELSE' 'IF' C=0.5 'THEN' M[K+1]:=0.0
'ELSE' 'IF' C=0.75 'THEN' M[K+1]:=M[K+1]*(-1/(K*T))
'ELSE' 'IF' C=1.0 'THEN' M[K+1]:=0.0;
OUTPUT(M[K+1]);
'END';

```

'COMMENT' THE NEXT SEGMENT CALCULATES THE FOURIER COEFFICIENTS AT SPECIAL FREQUENCIES (2J+1)*2PIE/NT THE RESULT IS DEPOSITED IN ARRAYS MSIN AND MCOS;

```

'BEGIN'
  RECYCLE:
    J:= -1;
    K:= 0.0;
    LOOP :
      K:=K+1;
      A:=((2*J)+1)*(K-1)/(G-1);
      A:=A-ENTIER(A);
      'IF' A>0.0 'AND' A <0.25 'THEN'
'BEGIN'
  MSIN[J+1]:=MSIN[J+1]+M[K]*SIN(2*PIE*A);
  MCOS[J+1]:=MCOS[J+1]+M[K]*COS(2*PIE*A);
'END'
  'ELSE' 'IF' A>0.25 'AND' A <0.5 'THEN'
'BEGIN'
  MSIN[J+1]:= MSIN[J+1]+ M[K]*SIN(PIE-2*PIE*A);
  MCOS[J+1]:=MCOS[J+1]+M[K]*(-1)*COS(PIE-2*PIE*A);
'END'
  'ELSE' 'IF' A>0.5 'AND' A <0.75 'THEN'
'BEGIN'
  MSIN[J+1]:= MSIN[J+1] +M[K]*SIN((-1)*(2*PIE*A -PIE));
  MCOS[J+1]:= MCOS[J+1] +M[K]*COS((-1)*(2*PIE*A -PIE));
'END'
  'ELSE' 'IF' A>0.75 'AND' A < 1.0 'THEN'
'BEGIN'
  MSIN[J+1]:= MSIN[J+1]+M[K]*SIN((-1)*(2*PIE-2*PIE*A));
  MCOS[J+1]:= MCOS[J+1]+M[K]*COS( (2*PIE-2*PIE*A));
'END'
  'ELSE' 'IF' A=0.0 'OR' A=1.0 'THEN'
    MCOS[J+1]:= MCOS[J+1] +M[K]
  'ELSE' 'IF' A=0.5 'THEN' MCOS[J+1]:=MCOS[J+1]+M[K]*(-1)
  'ELSE' 'IF' A=0.25 'THEN' MSIN[J+1]:=MSIN[J+1]+M[K]
  'ELSE' 'IF' A=0.75 'THEN' MSIN[J+1]:= MSIN[J+1]+M[K]*(-1)
'END';

  'IF' K = G 'THEN' 'GOTO' L6 'ELSE' 'GOTO' LOOP;
L6:  'IF' J=( E -1) 'THEN' 'GOTO' L7 'ELSE' 'GOTO' RECYCLE ;
L7:  'IF' H =1 'AND' AUTO 'THEN' 'GOTO' L8 'ELSE' 'IF' H=1
      'AND' NOT AUTO 'THEN' 'GOTO' L9 'ELSE' 'GOTO' L10;
L8:  'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
'BEGIN'
      MSIN[F] := MSIN[F];
      MCOS[F] := MCOS[F];
'END';
  'FOR' F:=1 'STEP' 1 'UNTIL' N 'DO'
'BEGIN'
  MSIN[F]:=0.0;
  MCOS[F]:=0.0;
'END';

  'IF' AUTO 'AND' CROSS 'THEN' 'GOTO' L11 'ELSE' 'GOTO'
L10;
L11: 'FOR' F:=1 'STEP' 1 'UNTIL' G 'DO'
      M[F] := RXY[(F+2)-1];
      'GOTO' RESTART;
L9 : 'GOTO' L13;
L10: 'COMMENT' THE FOLLOWING PROCEEDURE PLOTS THE AUTOCORRELOGRA

```


'BEGIN'

CK

5

'REAL' MAX, POWER;

'INTEGER' LITA;

'REAL' 'ARRAY' AB[1:5], BC[1:5] ;

'INTEGER' 'ARRAY' CD[1:17], DE[1:3], EF[1:22], FG[1:13],

GH[1:13], HI[1:24], JK[1:28], KL, MN, TU, VW[1:4], LM[1:5],

NO, OP, PQ, QR[1:7], RS[1:20], ST[1:25], UV[1:9];

'BEGIN'

CK

6

'PROCEDURE' OPENPLOT; 'EXTERNAL';

CK

7

'PROCEDURE' CLOSEPLOT; 'EXTERNAL';

CK

8

'PROCEDURE' HGPLOTT(X, Y, IC, L); 'VALUE' X, Y, IC, L; 'INTEGER' IC, L;

CK

9

'REAL' X, Y; 'EXTERNAL';

'PROCEDURE' STRARR(A, N, S); 'ARRAY' A; 'INTEGER' N;

CK

10

'STRING' S; 'EXTERNAL';

'PROCEDURE' HGPAXISVT(X, Y, BCD, NC, S, THETA, XMIN, DX, GAP, NH);

CK

11

'VALUE' X, Y, NC, S, THETA, XMIN, DX, GAP, NH; 'INTEGER' NC, NH; 'ARRAY'

BCD; 'REAL' X, Y, S, THETA, XMIN, DX, GAP; 'EXTERNAL';

'PROCEDURE' HGPLINET(X, Y, N, K); 'VALUE' N, K; 'ARRAY' X, Y;

CK

12

'INTEGER' N, K; 'EXTERNAL';

'PROCEDURE' HGPNUMBERT(X, Y, HT, FL, THETA, I, IP, IQ); 'VALUE' X, Y,

CK

13

HT, FL, THETA, I, IP, IQ; 'INTEGER' I, IP, IQ; 'REAL' X, Y, HT, FL, THETA;

'EXTERNAL';

'PROCEDURE' HGPSYMBLT(X, Y, HT, BCD, THETA, N); 'VALUE' X, Y, HT,

CK

14

THETA, N; 'INTEGER' N; 'ARRAY' BCD; 'REAL' X, Y, HT, THETA; 'EXTERNAL';

'PROCEDURE' PART 3;

CK

15

'BEGIN'

OPENPLOT ;

HGPLOTT(-3, 0, 10, 5, 0, 4);

MAX:=0.0;

'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

'BEGIN'

'IF' ABS(NCOS(J)) > MAX 'THEN' MAX:=ABS(NCOS(J)) 'ELSE' MAX:=MAX;

OUTPUT(NCOS(J));

'END';

CROSSNORM:=MAX;

NEWLINE(1);

OUTPUT(MAX);

'COMMENT' THE NEXT SEGMENT COMPUTES THE X AND Y COORDINATES TO
PLOT A HISTOGRAM;

A:=((2*PIE)/(G*T*SU));

D:=((ENTIER(((PIE/T)/SU)/8)+1)*8;

'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

'BEGIN'

AB[1]:=(((2*(J-1)+1)*A)-A)*4.5/D;

AB[3]:=(((2*(J-1)+1)*A)+A)*4.5/D;

AB[2]:= AB[1];

AB[4]:= AB[3];

```

BC[4] := 0.0;
HGPLINET(AB,BC,4,+1);
OUTPUT(NSINE[J]);
'END';

STRARR(CD,LITA,('FREQUENCYXRAD/SEC')));
HGPAXISVT(0.0,0.0,CD,-17.45,0.0,0.0,D/8,0.5625,-2);
STRARR(DE,LITA,('%%')));
HGPAXISVT(0.0,0.0,DE,0,3.5,90;0,0.0,0.5,1.25,+2);
STRARR(EF,LITA,('POWERXSPECTRALXDENSITY')));
HGPSYMBLT(-1.0,0.5,0.125,EF,90.0,22);
C:=(4.5+2*PIE)/D;
STRARR(FG,LITA,('CYCLES/SECOND')));
HGPAXISVT(0.0,-0.5,FG,-13.45,0.0,0.0,1.0,C,-1);
STRARR(GH,LITA,('CM,CM/RAD/SEC')));
HGPSYMBLT(-0.625,1.0,0.1,FG,90.0,13);
STRARR(HJ,LITA,('G.D.WHITEHEAD')));
HGPSYMBLT(3.5,8.875,0.1,GH,0.0,13);
STRARR(IK,LITA,('FOURIER/MILNERXTRANSFORM')));
HGPSYMBLT(-1.25,8.875,0.2,HJ,0.0,24);
STRARR(JK,LITA,('AUTOCORRELATIONXCOEFFICIENT')));
HGPSYMBLT(-1.0,4.75,0.125,JK,90.0,27);
HGPAXISVT(0.0,4.25,DE,0,3.5,90,0,-1,0.5,0.875,+1);
HGPAXISVT(0.0,6.0,DE,-0.45,0.0,0.0,(T*(N-1)*SU)/5,0.9,+1);
STRARR(KL,LITA,('TIME')));
HGPSYMBLT(4.5,6.125,0.1,KL,0.0,4);
STRARR(LM,LITA,('DELAY')));
HGPSYMBLT(4.5,5.75,0.1,LM,0.0,5);
STRARR(MN,LITA,('SECS')));
HGPSYMBLT(4.5,5.5,0.1,MN,0.0,4);
HGPLOTT(0.0,0.0,3.0);
'FOR' F:=1'STEP'1'UNTIL'N'DO'
M[F]:=(F/N*4.5)-(4.5/N);
'FOR' F:=1'STEP'1'UNTIL'N'DO'
RX[F]:=1.75*RX[F];
HGPLOTT(0.0,-6.0,0,4);
HGPLINET(M,RX,N,+1);
'FOR' F:=1'STEP'1'UNTIL'N'DO'
RXY[F]:=1.75*RXY[F];
STRARR(NO,LITA,('R(O)%%=')));
HGPSYMBLT(2.875,1.65,0.1,NO,0.0,7);
STRARR(OP,LITA,('R(INF)=')));
HGPSYMBLT(2.875,1.5,0.1,OP,0.0,7);
HGPNUMBERT(3.375,1.65,0.1,RX0,0.0,1.0,2,6);
HGPNUMBERT(3.375,1.50,0.1,AUTONORM,0.0,1.0,2,6);
STRARR(PQ,LITA,('POWER=')));
POWER:=0.0;
'FOR' F:=1'STEP'1'UNTIL'E'DO'
POWER:=POWER+NCOS[F];
HGPSYMBLT(2.875,-2.625,0.1,PQ,0.0,6);
HGPNUMBERT(3.125,-2.625,0.1,POWER,0.0,1.0,2,6);
STRARR(QR,LITA,('%%MAX=')));
HGPSYMBLT(2.875,-2.475,0.1,QR,0.0,6);
HGPNUMBERT(3.125,-2.475,0.1,MAX,0.0,1.0,2,6);
AB[1] := 5.25;
AB[2] := 5.25;
AB[3] := -1.75;

AB[4] := -1.75;

```



```

BC[3] := +3.25;
BC[4] := -7.25;
BC[5] := -7.25;
HGPLAINET(AB,BC,5,+1);
HGPLOTT(0,0,0.0,3,0);
L11:  HGPLOTT(-10.75,0.0,0,4);
HGPLAINET(M,RXY,N,+1);
HGPAxisVT(0.0,-1.75,DE,0,3.5,90.0,-1,0.5,0.875,+2);
HGPAxisVT(0.0,0.0,DE,-0.4,5,0.0,0.0,(T*(N-1)*SU)/5,0.9,+1);
STRARR(JK,LITA,('CROSSCORRELATION%COEFFICIENT'));
HGPSYMBLT(-1.0,-1.375,0.125,JK,90,0,28);
HGPSYMBLT(4.625,0.125,0.1,KL,0,0,4);
HGPSYMBLT(4.625,0.0,0.1,LM,0,0,5);
HGPSYMBLT(4.625,-0.125,0.1,MN,0,0,4);
MAX:=0,0;
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
'BEGIN'
'IF' ABS( MSIN[F] ) > MAX 'THEN' MAX :=ABS( MSIN[F] ) 'ELSE' MAX:=MAX;
'END';
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
'BEGIN'
'IF' ABS( MCOS[F] ) > MAX 'THEN' MAX:=ABS( MCOS[F] ) 'ELSE' MAX
:=MAX;
'END';
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
'BEGIN'
RX[F]:= (MCOS[F]/MAX)*2.5;
RXY[F]:= (MSIN[F]/MAX)*2.5;
'END';
HGPLOTT(0,0,0.0,3,0);
HGPLOTT(-1.75,4.625,0,4);
HGPLAINET(RX,RXY,E,+1);
HGPAxisVT(-2.5,0.0,DE,0,5.0,0.0,-1,0,0.2,0.5,1);
HGPAxisVT(0.0,-2.5,DE,0,5,0,90.0,-1,0,0.2,0.5,1);
STRARR(TU,LIYA,('REAL'));
STRARR(UV,LITA,('IMAGINARY'));
HGPSYMBLT(2.75,0.0,0.125,TU,0,0,4);
HGPSYMBLT(-0.5,2.7,0.125,UV,0,0,9);
STRARR(RS,LITA,('FOURIER/MILNER%GROSS'));
STRARR(ST,LITA,('SPECTRAL%DENSITY%ANALYSIS'));
HGPLOTT(0,0,0.0,3,0);
HGPLOTT(0.0,-1.875,0,4);
HGPSYMBLT(1.5,5.5,0.1,GH,0,0,13);
HGPSYMBLT(-2.875,5.5,0.125,RS,0,0,20);
HGPSYMBLT(-2.875,5.25,0.125,ST,0,0,25);
STRARR(VW,LITA,('MAX='));
HGPSYMBLT(-2.875,5.0,0.1,VW,0,0,4);
HGPNUMBERT(-2.275,5.0,0.1,MAX,0,0,1,0,6);
AB[1] := -3.5;
AB[2] := -3.5;
AB[3] := 3.5;
AB[4] := 3.5;
AB[5] := -3.5;
BC[1] := 6.0;
BC[2] := -4.5;
BC[3] := -4.5;
BC[4] := 6.0;
BC[5] := 6.0;
HGPLAINET(AB,BC,5,+1);

```

'PROCEDURE' PART 4;

OK 16

'BEGIN'

'COMMENT' THE NEXT SEGMENT PLOTS THE CO AND QUAD SPECTRUM;

```

HG PLOTT(0.0,0.0,3,0);
HG PLOTT(-10.0,0.0,0.4);
  HG P SYMBLT(-2.875,5.5,0.125,RS,0.0,20);
  HG P SYMBLT(-2.875,5.25,0.125,ST,0.0,25);
  HG P SYMBLT(1.5,5.5,0.1,GH,0.0,13);
  HG P LINET(AB,BC,5,+1);
HG PLOTT(0.0,0.0,3,0);
HG PLOTT(1.75,2.625,0,4);
  'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

```

'BEGIN'

```

  AB[1]:=(((2*(J-1)+1)*A)-A)*4.5/D;
  AB[3]:=(((2*(J-1)+1)*A)+A)*4.5/D;
  AB[2]:= AB[1];
  AB[4]:= AB[3];
  BC[1]:= 0.0;
  BC[2]:=MSIN[J]*2.5/MAX;
  BC[3]:= BC[2];
  BC[4]:= 0.0;
  HG P LINET(AB,BC,4,+1);

```

'END';

```

  HG P AXISVT(0.0,-1.25,CD,-17,4.5,0.0,0.0,0/8,0.5625,#2);
  HG P AXISVT(0.0,-1.25,EF,13,4.5,0.0,0.0,1.0,0,-1);
  HG P AXISVT(0.0,-1.25,DE,0,3.75,90.0,-0.5,0.5,1.25,+2);
  STRARK(JK,LITA,('QUADXSPECTRUM'));
  HG P SYMBLT(-1.0,-0.25,0.125,JK,90.0,13);
  HG P SYMBLT(-0.625,-0.25,0.1,FG,90.0,13);
  HG PLOTT(0.0,0.0,3,0);
  HG PLOTT(0.0,-4.375,0,4);
  'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'

```

'BEGIN'

```

  AB[1]:=(((2*(J-1)+1)*A)-A)*4.5/D;
  AB[3]:=(((2*(J-1)+1)*A)+A)*4.5/D;
  AB[2]:= AB[1];
  AB[4]:= AB[3];
  BC[1]:= 0.0;
  BC[2]:=MCOS[J]*2.5/MAX;
  BC[3]:= BC[2];
  BC[4]:= 0.0;
  HG P LINET(AB,BC,4,+1);

```

'END';

```

  HG P AXISVT(0.0,-1.25,CD,-17,4.5,0.0,0.0,0/8,0.5625,#2);
  HG P AXISVT(0.0,-1.25,EF,13,4.5,0.0,0.0,1.0,0,-1);
  HG P AXISVT(0.0,-1.25,DE,0,3.75,90.0,-0.5,0.5,1.25,+2);
  STRARK(JK,LITA,('COXSPECTRUM'));
  HG P SYMBLT(-1.0,0.0,0.125,JK,90.0,11);
  HG P SYMBLT(-0.625,-0.25,0.1,FG,90.0,13);

```

'END' OF PART 4;

'PROCEDURE' PART 5;

OK 17

'BEGIN'

'COMMENT' THE NEXT SEGMENT COMPUTES AND PLOTS THE GAIN & PHASE;

```

  'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
  RX[F]:=20.0+0.43429448*(LN(SQRT((MSIN[F]*MSIN[F])
  +(MCOS[F]*MCOS[F]))/
  SQRT((NSIN[F]*NSIN[F])+(NCOS[F]*NCOS[F]))));
  'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'

```



```

RX(F):=RX[F]*1.875/20;
HG PLOTT(0,0,0,0,3,0);
HG PLOTT(10,0,14.25,0,4);
  'FOR' J:=1 'STEP' 1 'UNTIL' E 'DO'
'BEGIN'
  AB[1]:=((2*(J-1)+1)*A)-A)*4.5/D;
  AB[S]:=((2*(J-1)+1)*A)+A)*4.5/D;
  AB[2] := AB[1];
  AB[4] := AB[3];
  BC[1] := 0.0;
  BC[2]:=RX[J];
  BC[3] := BC[2];
  BC[4] := 0.0;
  HG PLINET(AB,BC,4,+1);
'END';

HG PAXISVT(0,0,0,0,DF,-0,4.5,0,0,0,0,D/8,0.5625,-1);
HG PAXISVT(0,0,-1.875,CD,-17,4.5,0,0,0,0,D/8,0.5625,-2);
HG PAXISVT(0,0,-1.875,EF,13,4.5,0,0,0,0,1,0,C,-1);
HG PAXISVT(0,0,-1.875,DE,0,3.75,90,0,-20,10,0.9375,-2);
STRARR(JK,LITA,('GAIN=DECIBELS'));
HG PSYMBLT(-1,0,-0.875,0.125,JK,90,0,13);
HG PLOTT(0,0,0,0,3,0);
HG PLOTT(0,0,4.375,0,4);
HG PNUMBERT(2,25,1.1,0.1,TD,0,0,0,2,3);
STRARR(JK,LITA,('M%SECS. '));
HG PSYMBLT(3,0,1.1,0.1,JK,0,0,7);
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
RX[F]:=-360*TD/((G*T*SU)/(2*(F-1)+1));
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
'BEGIN'
  S1:=SIGN(MCOS[F]);
  S2:=SIGN(MSIN[F]);
  T:=ABS(ARCTAN(MSIN[F]/MCOS[F])+57.29578);
  'IF' S1=1 'AND' S2=-1 'AND' F<17 'THEN' RX[F]:=RX[F]+Y
'ELSE'
  'IF' S1=0 'AND' S2=0 'THEN' RX[F]:=RX[F]+0,0
  'ELSE' 'IF' S1=1 'AND' S2=1 'THEN' RX[F]:=RX[F]+T
  'ELSE' 'IF' S1=0 'AND' S2=1 'THEN' RX[F]:=RX[F]-90
  'ELSE' 'IF' S1=0 'AND' S2=-1 'THEN' RX[F]:=RX[F]+270
  'ELSE' 'IF' S1=1 'AND' S2=0 'THEN' RX[F]:=RX[F]+0,0
  'ELSE' 'IF' S1=-1 'AND' S2=0 'THEN' RX[F]:=RX[F]-180
  'ELSE' 'IF' S1=-1 'AND' S2=1 'THEN' RX[F]:=RX[F]-(180-T)
  'ELSE' 'IF' S1=1 'AND' S2=-1 'THEN' RX[F]:=RX[F]-(360-T)
  'ELSE' 'IF' S1=-1 'AND' S2=-1 'THEN' RX[F]:=RX[F]-(180+T);
  'IF' RX[F]<=-360 'THEN' RX[F]:=-360;
'END';
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO' RX[F]:=RX[F]*1.875/360;
  'FOR' F:=0 'STEP' 1 'UNTIL' (E-1) 'DO'
  RXY[F+1]:=((2*(F)+1)*A)*4.5/D;
  HG PLINET(RXY,RX,E,+1);
  HG PAXISVT(0,0,0,0,DF,-0,4.5,0,0,0,0,D/8,0.5625,-1);
  HG PAXISVT(0,0,-1.875,CD,-17,4.5,0,0,0,0,D/8,0.5625,-2);
  HG PAXISVT(0,0,-1.875,EF,13,4.5,0,0,0,0,1,0,C,-1);
  HG PAXISVT(0,0,-1.875,DE,0,3.75,90,0,-360,180,0.9375,-2);
  STRARR(JK,LITA,('PHASE%ANGLE%=DEGREES'));
  HG PSYMBLT(-1,0,-1,0,0.125,JK,90,0,20);
  STRARR(JK,LITA,('AGGREGATE%TIME%DELAY'));
  HG PSYMBLT(2,25,1,25,0.1,JK,0,0,20);
  HG PLOTT(0,0,0,0,3,0);
  HG PLOTT(-1,75,-1,75,0,4);

```

```
HGPSYMBL(-2.875,5.25,0.125,ST,0.0,25);
HGPSYMBL(1.5,5.5,0.1,GH,0.0,13);
```

```
AB[1] := -3.5;
AB[2] := -3.5;
AB[3] := 3.5;
AB[4] := 3.5;
AB[5] := -3.5;
```

```
BC[1] := 6.0;
BC[2] := 4.5;
BC[3] := 4.5;
BC[4] := 6.0;
BC[5] := 6.0;
```

```
HUPLINET(AB,BC,5,*1);
```

```
L12: CLOSEPLOT;
```

```
SELECT OUTPUT(4);
```

```
RUNOUT;
```

```
RUNOUT;
```

```
NEWLINE(1);
```

```
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
```

```
'BEGIN'
```

```
PRINT(NCOS[F],2,5);
```

```
NEWLINE(1);
```

```
'END';
```

```
RUNOUT;
```

```
RUNOUT;
```

```
NEWLINE(1);
```

```
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
```

```
'BEGIN'
```

```
PRINT(NSIN[F],2,5);
```

```
NEWLINE(1);
```

```
'END';
```

```
RUNOUT;
```

```
RUNOUT;
```

```
NEWLINE(1);
```

```
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
```

```
'BEGIN'
```

```
PRINT(MCOS[F],2,5);
```

```
NEWLINE(1);
```

```
'END';
```

```
RUNOUT;
```

```
RUNOUT;
```

```
NEWLINE(1);
```

```
'FOR' F:=1 'STEP' 1 'UNTIL' E 'DO'
```

```
'BEGIN'
```

```
PRINT(MSIN[F],2,5);
```

```
NEWLINE(1);
```

```
'END';
```

```
RUNOUT;
```

```
RUNOUT;
```

```
'END' OF PART 5;
```

```
PART 1;
```

```
PART 2 (L13);
```

```
PART 3;
```

```
PART 4;
```

```
PART 5;
```

```
'END';
```

```
'END';
```

```
L13: 'END';
```

```
'END';
```


JOB TWO FIVE, ASSTF215, WHITHEAD, STREAM 3 RUN AT ASTON UNIV.
BULLETIN TWO FIVE , 3,300
VOLUME 5000
ALGOLUB

JOB COST NO	202	DATE	14/12/70
JOB NAME	TWO FIVE	START TIME	11/31/27
USER NAME	WHITHEAD	END TIME	11/44/26
PERIPHERALS USED:	30		

TOTAL MILL TIME	274
INPUT RECORDS	1561
OUTPUT RECORDS	1489
MAX. CORE SIZE	16000

25/02/71

COMPILED BY XALT MK. 3

'SEND TO' (ED,ICLA-DEFAULT(0),.PROGRAM)

'WORK' (ED,WORK FILE (0))

'BEGIN'

'REAL' A,D,NORM,RX0,C,PIE,AUTONORM,CROSSNORM,T,TD,P;
'INTEGER' K,N,J,F,H,Z,E,SU,S1,S2,G;

T:=0.02;

PIE := 3.14159265358979323846;

SU:=4;

SELECT INPUT(3);

N:=READ;

H:=READ;

C:=READ;

C:=60.0;

G:=N*H;

'BEGIN'

'REAL' 'ARRAY' RX[1:G],RXY[1:G],M[1:G];

'BEGIN'

'REAL' MAX,POWER;

'INTEGER' LITA;

'REAL' 'ARRAY' AB[1:5],BC[1:5];

'INTEGER' 'ARRAY' CD[1:17],DE[1:3],FF[1:22],FG[1:13],

GH[1:13],HI[1:24],JK[1:28],KL,MM,TU,VW[1:4],LM[1:5],

NO,OP,PO,QR[1:7],RS[1:20],ST[1:25],UV[1:9];

'FOR' J:=1'STEP'1'UNTIL' H'DO' CD[J]:=READ;

'BEGIN'

'PROCEDURE' OPENPLOT;'EXTERNAL';

'PROCEDURE' CLOSEPLOT;'EXTERNAL';

'PROCEDURE' HGPLOT (X,Y,IC,L);'VALUE'X,Y,IC,L;'INTEGER'IC,L;
'REAL'X,Y;'EXTERNAL';

'PROCEDURE' STRARR(A,N,S);'ARRAY'A;'INTEGER' N;

'STRING' S;'EXTERNAL';

'PROCEDURE' HGRAXISV (X,Y,BCD,NC,S,THETA,XMIN,DX,GAP,NH);

'VALUE'Y,Y,NC,S,THETA,XMIN,DX,GAP,NH;'INTEGER'NC,NH;'ARRAY'
BCD;'REAL'X,Y,S,THETA,XMIN,DX,GAP;'EXTERNAL';

'PROCEDURE' HGPLINE (X,Y,N,K);'VALUE'N,K;'ARRAY'X,Y;

'INTEGER' N,K;'EXTERNAL';

'PROCEDURE' HGPNUMBER (X,Y,HT,FL,THETA,I,IP,IQ);'VALUE'X,Y,
HT,FL,THETA,I,IP,IQ;'INTEGER'I,IP,IQ;'REAL'X,Y,HT,FL,THETA;

'EXTERNAL';

'PROCEDURE' HGPSYMBL (X,Y,HT,BCD,THETA,N);'VALUE'X,Y,HT,
THETA,N;'INTEGER'N;'ARRAY'BCD;'REAL'X,Y,HT,THETA;'EXTERNAL';

'FOR' F:=1'STEP'1'UNTIL'G'DO'

RX[F]:=READ;

'FOR' J:=1'STEP'1'UNTIL' H'DO'

'BEGIN'

'FOR' F:=1'STEP'1'UNTIL' N'DO'

RX[F+(N*(J-1))]:=RX[F+(N*(J-1))+CD[J];

'END';

A:=0.0;

'BEGIN'

'FOR' F:=1'STEP'1'UNTIL'G'DO'

'END';

'IF' ABS(RX[F])> A 'THEN' A:=ABS(RX[F]);

'FOR' F:=1'STEP'1'UNTIL'G'DO'

RX[F]:=PX[F]*1.56/A;

AUTONORM:=(4.3/(H-1));

NORM:=(4.3/(H-1))/(SIN(C/57.29578)/COS(C/57.29578));

```

HG PLOT(-6,0,12.0,0,4);
'FOR' J:=1'STEP'1'UNTIL' N'DO'
'BEGIN'
Z:=0,0;
'FOR' F:=1'STEP'1'UNTIL' H'DO'
'BEGIN'
M[F]:=AUTONORM*(F-1)+1,66-(ABS(RX[F+Z+J-1]));
Z:=Z+N-1;
'END'
'FOR' F:=1'STEP'1'UNTIL' H'DO'
RXY[F]:=((F-1)*NORM)+((J-1)*(3,0/(N-1)));
HG PLINE(RXY,M,H,+1);
'END'
'FOR' J:=1'STEP'1'UNTIL' H'DO'
'BEGIN'
'FOR' F:=1'STEP'1'UNTIL' N'DO'
M[F]:=AUTONORM+(J-1)+1,66-(ABS(RX[F+(N*(J-1))]));
'FOR' F:=1'STEP'1'UNTIL' N'DO'
RXY[F]:=(F-1)*3,0/(N-1)+NORM*(J-1);
HG PLINE(RXY,M,N,+1);
'END'
AB[1]:=0,0;
AB[2]:=0,0;
AB[3]:=3,0;
AB[4]:=4,3/(SIN(C/57,29578)/(COS(C/57,29578)))+3,0;
BC[1]:=1,66;
BC[2]:=0,0;
BC[3]:=0,0;
BC[4]:=4,3;
HG PLINE (AB,BC,4,+1);
D:=(CENTIER(((PIF/T)/SU)/8)+1)*8;
A:=(3,0*2*PIE)/D;
STRARR(CD,LITA,('FREQUENCY%RAD/SEC')));
HG PAXISV (0,0,0,0,0,-17,3,0,0,0,0,0/8,0,3750,-1);
STRARR(EF,LITA,('POWER%SPECTRAL%DENSITY')));
HG PSYMBL (-0,625,0,5,0,1,EF,90,0,22);
STRARR(CF,LITA,('CYCLES/SECOND')));
HG PAXISV (0,0,-0,5,EF,-13,4,5,0,0,0,0,1,0,A,-1);
STRARR(CH,LITA,('G.D.WHITEHEAD')));
HG PSYMBL (R,0,7,95,0,1,GH,0,0,13);
AR[1]:=-1,25;
AR[2]:=5,65;
AR[3]:=0,65;
AR[4]:=-1,25;
AR[5]:=-1,25;
BC[1]:=-2,0;
BC[2]:=-2,0;
BC[3]:=8,45;
BC[4]:=8,45;
BC[5]:=-2,0;
HG PLINE(AB,BC,5,+1);
AR[1]:=-1,0;
AR[2]:=6,25;
AR[3]:=6,25;
AR[4]:=-1,9;
AR[5]:=-1,9;
BC[1]:=-2,65;
BC[2]:=-2,65;
BC[3]:=-2,65;

```


BC[5]:=-2.65;

HGPLINE(AB,BC,3,+1);
 HGPOINT(0,0,0.0,3,0);
 HGPOINT(-12.5,0,0,4);
 'FOR' J:=1'STEP'1'UNTIL' H 'DO'

'BEGIN'

'FOR' F:=1'STEP'1'UNTIL' N 'DO'
 M[F]:=ABS(RX[F+(N*(J-1))])+AUTONORM*(J-1);
 'FOR' F:=1'STEP'1'UNTIL' N 'DO'
 RXY[F]:=(F-1)*3.0/(N-1)+NORM*(J-1);
 HGPLINE(RXY,M,N,+1);

'END'

'FOR' J:=1'STEP'1'UNTIL' N'DO'

'BEGIN'

Z:=0,0;
 'FOR' F:=1'STEP'1'UNTIL' H 'DO'

'BEGIN'

H[F]:=AUTONORM*(F-1)+ (ABS(RX[F+Z+J-1]));
 Z:=Z+N*1;

'END'

'FOR' F:=1'STEP'1'UNTIL' H 'DO'
 RXY[F]:=((F-1)*NORM)+(J-1)*(3.0/(N-1));
 HGPLINE(RXY,H,H,+1);

'END'

AB[1]:=0,0;
 AB[2]:=0,0;
 AB[3]:=3,0;
 AB[4]:=4.3/(SIN(C/57.29578)/(COS(C/57.29578)))+3,0;
 BC[1]:=1,66;

BC[2]:=0,0;

BC[3]:=0,0;

BC[4]:=4,3;

HGPLINE(AB,BC,4,+1);
 STRARR(CD,LITA,('FREQUENCYXRAD/SEC'));
 HGPAXISV(0,0,0,0,CD,-17,3,0,0,0,0,0,0,0,0,0,8,0.3750,-1);
 STRARR(EF,LITA,('POWERXSPECTRALXDENSITY'));
 HGPSYMBL(-0.625,0.5,0.1,EF,90.0,22);
 STRARR(EF,LITA,('CYCLES/SECOND'));
 HGPAXISV(0,0,-0.5,EF,-13,4.5,0,0,0,0,1,0,A,-1);
 STRARR(GH,LITA,('G.D.WHITHEAD'));

HGPSYMBL(3,9,7,95,0.1,GH,0,0,13);

AB[1]:=-1.25;

AB[2]:=5.65;

AB[3]:=5.65;

AB[4]:=-1.25;

AB[5]:=-1.25;

BC[1]:=-2,0;

BC[2]:=-2,0;

BC[3]:=8,45;

BC[4]:=8,45;

BC[5]:=-2,0;

HGPLINE(AB,BC,5,+1);

AB[1]:=-1,9;

AB[2]:=6,25;

AB[3]:=6,25;

AB[4]:=-1,9;

AB[5]:=-1,9;

BC[1]:=-2,65;

BC[2]:=-2,65;

BC[4]:=9.1;
BC[5]:=-2.65;

HGPLINE(AB,BC,5,+1);
HG PLOT(0,0,0,0,3,0);
HG PLOT(0,0,13.0,0,4);
!FOR' J:=1'STEP'1'UNTIL' H 'DO'

!BEGIN'

!FOR' F:=1'STEP' 1 'UNTIL' N 'DO'
M[F]:=ABS(RX[F+(N*(J-1))])+AUTONORM*(J-1);
!FOR' F:=1'STEP' 1 'UNTIL' N 'DO'
RXY[F]:=(F-1)*3.0/(N-1)+NORM*(J-1);
HGPLINE(RXY,M,N,+1);

!END';

AB[1]:=0,0;
AB[2]:=0,0;
AB[3]:=3,0;
AB[4]:=4.3/(SIN(C/57.29578)/(COS(C/57.29578)))+3,0;
BC[1]:=1.66;

BC[2]:=0,0;
BC[3]:=0,0;
BC[4]:=4,3;

HG PLINE (AB,BC,4,+1);

AB[1]:=-1.25;
AB[2]:=5.65;
AB[3]:=5.65;
AB[4]:=-1.25;
AB[5]:=-1.25;
BC[1]:=-2,0;
BC[2]:=-2,0;
BC[3]:=8,45;
BC[4]:=8,45;
BC[5]:=-2,0;

HG PLINE(AB,BC,5,+1);

AB[1]:=-1.9;
AB[2]:=6.25;
AB[3]:=6,25;
AB[4]:=-1.9;
AB[5]:=-1.9;
BC[1]:=-2.65;
BC[2]:=-2.65;
BC[3]:=9,1;
BC[4]:=9,1;
BC[5]:=-2.65;

HG PLINE(AB,BC,5,+1);
CLOSEPLOT;

!END';
!END';
!END';
!END';

LENGTH 2233

ED

47

EC

JOB DISPLAY,ASSTF215,WHITHEAD,STREAM 3 RUN AT ASTON UNIV.
BULLETIN DISPLAY ,3,300
ALGOL ,,,,BINARYSTORE F215ISD,AXXXJ
STOREDPRG F215ISD

JOB COST NP	196	DATE	25/02/71
JOB NAME	DISPLAY	START TIME	19/33/44
USER NAME	WHITHEAD	END TIME	19/46/38
PERIPHERALS USED;			

TOTAL MILL TIME	120
INPUT RECORDS	689
OUTPUT RECORDS	308
MAX. CORE SIZE	18752

-473-
PHASE PLANE ANALYSIS PROGRAM

16/07/70

COMPILED BY XALE MK. 4B

'DUMP ON' (ED, PROGRAM TEMP(6), .NEWFILE)

'RUN'

'LIBRARY' (ED, SUBGROUPSRA3)

'LIBRARY' (ED, SUBGROUPGRAH)

'INPUT' 3=TR0

'OUTPUT' 0=LPO

'BEGIN' 'REAL' A, B, C, D, VEL1, VEL2, VEL3, ACCEL1, ACCEL2, AMPAV1, AMPAV2,
 AMPAV3, AMDIF1, AMDIF2, ANGLE1, ANGLE2, ANGDIF, ANGVEL, SAMPINT,
 VOLTRANGE, ADDRESS, HISTOGRAM, W, X, FORG1, DYNAMRANGE;
 'INTEGER' SAMPLESIZE, I, J, K, L, COUNTER, E, F, G, H, R, S, N, P, Q, Y, Z, LITA,
 LITB, LITC, LITD, ONE, SGN1, SGN2, SGN3, SGN4, SGN5, SIGN1, SIGN2,
 ALPHA, QUAD1, QUAD2, QUAD3, QUAD4;

'BOOLEAN' BOOL1, BOOL2, BOOL3, BOOL4, BOOL5;

A:=B:=C:=D:=ADDRESS:=HISTOGRAM:=0;

SELECT OUTPUT (0);

SELECT INPUT (3);

SAMPLESIZE := READ;

I := READ;

J := READ;

K := READ;

L := READ;

SAMPINT := READ;

VOLTRANGE := READ;

DYNAMRANGE := READ;

QUAD3 := READ;

G:=2*J+1;

'BEGIN' 'INTEGER' 'ARRAY' AMP[I:J], DEG[0:360], ANG[K:L], REF[1:20],
 FGE[1:14], GH, HJ, JK[1:2], BCD[1:4], LM[1:14], MN, NO[1:10],
 OR[1:7], RS[1:11], UV[1:25], ABC[1:17], DEF[1:29], XYZA
 [1:22], XYZCE[1:9], KL[1:16];

'REAL' 'ARRAY' PMA[1:SAMPLESIZE], LEVGNA[1:SAMPLESIZE],
 AB, BC, CD, DE, UVM, XYZ, ST, TU[1:5], GHI[1:12], OP, PQ[1:36],
 XY, YZ, JKL, MNO, PQR, STU[1:6];

'FOR' Y:=1 'STEP' 1 'UNTIL' J 'DO' AMP[Y]:=0;

'FOR' Y:=1 'STEP' 1 'UNTIL' SAMPLESIZE 'DO' PMA[Y]:=0;

'FOR' Z:=0 'STEP' 1 'UNTIL' 360 'DO' DEG[Z]:=0;

'FOR' Z:=K 'STEP' 1 'UNTIL' L 'DO' ANG[Z]:=0;

'FOR' Z:=1 'STEP' 1 'UNTIL' SAMPLESIZE 'DO' LEVGNA[Z]:=0;

SELECT INPUT (3);

F:=READ;

A:=READ;

F:=READ;

B:=READ;

F:=READ;

C:=READ;

'FOR' W:=A, B, C 'DO'

'BEGIN' HISTOGRAM := (W/VOLTRANGE)*J;

HISTOGRAM := ENTIER(HISTOGRAM);

'IF' HISTOGRAM > J-1 'THEN' F:=J 'ELSE' 'IF' HISTOGRAM < I+1 'THEN'

F:=I 'ELSE' F:=HISTOGRAM;

AMP[F]:=AMP[F]+1;

'END';

N:=3;

VEL1:=(A-B);

```

AMPV2:= (B+C)/2;
AMDIF1:= AMPV1-AMPV2;
SGN1 := SIGN(ACCEL1);
SGN2 := SIGN(AMDIF1);
'IF' SGN1=0 'AND' SGN2=0 'THEN' 'GOTO' L2 'ELSE' 'IF' SGN1=0 'THEN'
'GOTO' L5 'ELSE' 'IF' SGN2=0 'THEN' 'GOTO' L4;
'IF' SGN1=1 'AND' SGN2=1 'THEN' QUAD1:=1 'ELSE' 'IF' SGN1=1 'AND'
SGN2=-1 'THEN' QUAD1:=2 'ELSE' 'IF' SGN1=-1 'AND' SGN2=-1 'THEN'
QUAD1:=3 'ELSE' 'IF' SGN1=-1 'AND' SGN2=1 'THEN' QUAD1:=4;
ANGLE1 := ABS(ARCTAN(ACCEL1/AMDIF1)* 57.29578);
'IF' QUAD1=2 'THEN' ANGLE1:= 180-ANGLE1 'ELSE' 'IF' QUAD1=3
'THEN' ANGLE1:= 180+ANGLE1 'ELSE' 'IF' QUAD1=4 'THEN' ANGLE1:=
360 - ANGLE1 'ELSE' ANGLE1 := ANGLE1 ;
'GOTO' L5;
L2: QUAD1:=1;
ANGLE1:=0;
'GOTO' L5;
L3: 'IF' SGN2= -1 'THEN' 'BEGIN' ANGLE1:= 180; QUAD1:=2; 'END'
'ELSE'
'IF' SGN2=+1 'THEN' 'BEGIN' ANGLE1:=0; QUAD1:=1; 'END';
'GOTO' L5;
L4: 'IF' SGN1=-1 'THEN' 'BEGIN' ANGLE1:=270; QUAD1:=3; 'END' 'ELSE'
'IF' SGN1=+1 'THEN' 'BEGIN' ANGLE1:=90; QUAD1:=1; 'END';
'GOTO' L5;
L5: F:= ENTIER(ANGLE1);
DEG[F]:=DEG[F]+1;

```

RECYCLE: SELECT INPUT (3) ;
F:=READ;

```

C:= READ;
VEL1:= (A-B);
VEL2:= (B-C);
ACCEL2:= VEL1-VEL2;
AMPV1:= (A+B)/2;
AMPV2:= (B+C)/2;
AMDIF2:= AMPV1-AMPV2;
SGN3 := SIGN(ACCEL2);
SGN4 := SIGN(AMDIF2);
'IF' SGN3=0 'AND' SGN4=0 'THEN' 'GOTO' L6 'ELSE' 'IF' SGN3=0 'THEN'
'GOTO' L7 'ELSE' 'IF' SGN4=0 'THEN' 'GOTO' L8;
'IF' SGN3=1 'AND' SGN4=1 'THEN' QUAD2:=1 'ELSE' 'IF' SGN3=1 'AND'
SGN4=-1 'THEN' QUAD2:=2 'ELSE' 'IF' SGN3=-1 'AND' SGN4=-1
'THEN' QUAD2:=3 'ELSE' 'IF' SGN3=-1 'AND' SGN4=+1 'THEN'
QUAD2 :=4;
ANGLE2 := ABS(ARCTAN(ACCEL2/AMDIF2)*57.29578);
'IF' QUAD2 =2 'THEN' ANGLE2:= 180-ANGLE2 'ELSE' 'IF' QUAD2=3
'THEN' ANGLE2 := 180+ANGLE2 'ELSE' 'IF' QUAD2 =4 'THEN'
ANGLE2 :=360-ANGLE2 'ELSE' ANGLE2 := ANGLE2 ;
'GOTO' L9;
L6: QUAD2:=1;
ANGLE2:=0;
'GOTO' L9;
L7: 'IF' SGN4=-1 'THEN' 'BEGIN' ANGLE2:= 180; QUAD2:=2; 'END'
'ELSE' 'IF' SGN4=1 'AND' QUAD2=1 'OR' QUAD2=2 'THEN' 'BEGIN'
ANGLE2:=0; QUAD2:=1; 'END' 'ELSE' 'IF' SGN4=1 'AND' QUAD2=3 'OR'
QUAD2=4 'THEN' 'BEGIN' ANGLE2 := 360; QUAD2 :=4; 'END';

```

```

      'GOTO' L9;
L9: 'IF' QUAD1=1 'AND' QUAD2=4 'OR' QUAD1=4 'AND' QUAD2=1
      'THEN' 'GOTO' L10 'ELSE' 'GOTO' L11;
L10: 'IF' QUAD1=4 'THEN' ANGDIFF:=(360-ANGLE1)-ANGLE2
      'ELSE' 'IF' QUAD2=4 'THEN' ANGDIFF:=(ANGLE1-(360-ANGLE2));
      'GOTO' L12;
      L11: ANGDIFF := ANGLE1-ANGLE2;
      L12: ANGVEL := (ANGDIFF/SAMPINT);
           HISTOGRAM := (C/VOLTRANGE)*J;
           HISTOGRAM := ENTIER(HISTOGRAM);
           'IF' HISTOGRAM > J-1 'THEN' F:=J 'ELSE' 'IF' HISTOGRAM < I+1
           'THEN' F:=I 'ELSE' F:=HISTOGRAM;
           AMPE[F]:= AMPE[F]+1;
           ADDRESS:=(ANGVEL/720)*L;
           ADDRESS := ENTIER(ADDRESS);
           'IF' ADDRESS > L-1 'THEN' G:=L 'ELSE' 'IF' ADDRESS < K+1 'THEN'
           G:=K 'ELSE' G:= ADDRESS;
           ANG[G]:= ANG[G]+1;
           Z := F;
           S := G;
           F:= ENTIER(ANGLE2);
           DEG[F] := DEG[F]+1;
           N := N+1;
           PHA[N-3]:= (C/VOLTRANGE)* 3.0;
           LEVNA[N-3] := (ANGDIFF/720)* 3.0;
           'IF' N 'LE' SAMPLESIZE 'THEN' 'GOTO' REARRANGE 'ELSE' 'GOTO'
           FINISH;
REARRANGE: A:=B;
           S:=C;
           QUAD1:=QUAD2;
           ANGLE1:= ANGLE2;
           'GOTO' RECYCLE;
FINISH:
           'FOR' F:=0 'STEP' 1 'UNTIL' 360 'DO'
           'BEGIN' OP[F+1]:=(DEG[F]/(N-2))* 6.0;
                   PQ[F+1]:= -3.0 +(F*6/360);
           'END';
           G:= (L*2)+1;
           'FOR' F:=K 'STEP' 1 'UNTIL' L 'DO'
           'BEGIN' P:= ABS(L + F);
                   XY[P+1] := (ANG[F]/(N-3))* 6.0;
OUTPUT(ANG[F]);
                   YZ[P+1]:= -3.0 +ABS(P*(3/K));
           'END';
NEWLINE(1);
           'FOR' F:= I 'STEP' 1 'UNTIL' J 'DO'
           'BEGIN' P:= ABS((J)+(F));
                   JKL[P+1]:= (AMPE[F]/N)* 6.0;
OUTPUT(AMPE[F]);
                   MNO[P+1]:= -3.0+(ABS(P*(3/J)));
           'END';
'BEGIN'
'PROCEDURE' OPENPLOT; 'EXTERNAL';
'PROCEDURE' CLOSEPLOT; 'EXTERNAL';
'PROCEDURE' HGLOT(X,Y,IC,L); 'VALUE' X,Y,IC,L; 'INTEGER' IC,L;
'REAL' X,Y; 'EXTERNAL';
'PROCEDURE' ... 'INTEGER' N;

```



```

BCD;'REAL'X,Y,S,THETA,XMIN,DX,GAP;'EXTERNAL';
'PROCEDURE' HGPLINET(X,Y,N,K);'VALUE'N,K;'ARRAY'X,Y;
'INTEGER' N,K;'EXTERNAL';
'PROCEDURE' HGPNUMBERT(X,Y,HT,FL,THETA,I,IP,IQ);'VALUE'X,Y,
HT,FL,THETA,I,IP,IQ;'INTEGER'I,IP,IQ;'REAL'X,Y,HT,FL,THETA;
'EXTERNAL';
'PROCEDURE' HGPSYMBLT(X,Y,HT,BCD,THETA,N);'VALUE'X,Y,HT,
THETA,N;'INTEGER'N;'ARRAY'BCD;'REAL'X,Y,HT,THETA;'EXTERNAL';
OPENPLOT;
HGPLOTT(-5.0,8.5,0,4);

```

'BEGIN'

```

'COMMENT' SERIES ANGAMP PLOT;
HGPLINET(PMA,LEVGN,SAMPLESIZE=3,+1);
STRARR(BCD,LITA,('%%'));
HGPAXISVT(-3.0,3.0,BCD,0,6.0,0.0,-10,1.0,0.3,-1);
HGPAXISVT(0.0,-3.0,BCD,3,6.0,90.0,-10,1.0,0.3,-1);
AB[1] := -3.0;
AB[2] := 3.0;
AB[3] := 3.0;
AB[4] := -3.0;
AB[5] := -3.0;
BC[1] := -3.0;
BC[2] := -3.0;
BC[3] := 3.0;
BC[4] := 3.0;
BC[5] := -3.0;
HGPLINET(AB,BC,5,+1);
CD[1] := -4.125;
CD[2] := 4.125;
CD[3] := 4.125;
CD[4] := -4.125;
CD[5] := -4.125;
DE[1] := -5.5;
DE[2] := -5.5;
DE[3] := 6.25;
DE[4] := 6.25;
DE[5] := -5.5;
HGPLINET(CD,DE,5,+1);
STRARR(EF,LITA,('ANGAMP% DISTRIBUTION'));
HGPSYMBLT(-3.0,5.125,0.25,EF,0.0,19);
STRARR(FG,LITA,('G.D.WHITEHEAD'));
HGPSYMBLT(2.5,5.625,0.125,FG,0.0,13);
STRARR(GH,LITA,('N='));
HGPSYMBLT(2.5,4.5,0.125,GH,0.0,2);
HGPNUMBERT(2.75,4.5,0.125,SAMPLESIZE,0.0,1,4,0);
STRARR(HJ,LITA,('T='));
HGPSYMBLT(2.5,4.25,0.125,HJ,0.0,2);
STRARR(JK,LITA,('V='));
HGPSYMBLT(2.5,4.0,0.125,JK,0.0,2);
HGPNUMBERT(2.75,4.0,0.125,VOLTRANGE,0,0,1,5,0);
STRARR(KL,LITA,('ANGULAR%VELOCITY'));
HGPSYMBLT(-0.75,3.5,0.125,KL,0.0,16);
STRARR(LM,LITA,('DEG % SEC % * 36/T'));
HGPSYMBLT(-0.75,3.25,0.125,LM,0.0,13);
STRARR(MN,LITA,('AMPLITUDE'));
HGPSYMBLT(3.125,0.125,0.125,MN,0.0,9);
STRARR(NO,LITA,('VOLTS*V/10'));
HGPSYMBLT(3.125,0.125,0.125,NO,0.0,10);

```

```
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
WRITETEXT('('('('('0')' SUCCESS1')')');
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
'COMMENT'                                DEGREE PLOT ;
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
'BEGIN'
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
HGPLOTT(-10.0,2.0,0,4);
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
HGPLINET(PQ,OP,361,+1);
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
STRARR(QR,LITA,'('DEGREES')');
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
HGPAXISVT(-3.0,0.0,QR,7,6.0,0.0,0.0,45,0.75,-1);
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
STRARR(RS,LITA,'('PROBABILITY')');
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
HGPAXISVT(0.0,0.0,RS,11,6.0,90.0,0.0,1.0,0.6,-1);
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
ST[1] := -4.125;
F:=F+1;
NEWLINE(1);
PRINT(F,2,0);
ST[2] := 4.125;
F:=F+1;
NEWLINE(1);
PRINT(F,2,0);
ST[3] := 4.125;
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
ST[4] := -4.125;
F:=F+1;
PRINT(F,2,0);
NEWLINE(1);
ST[5] := -4.125;
F:=F+1;
```

```

      F:=F+1;
PRINT(F,2,0);
  NEWLINE(1);
  TU[2] := -2.5;
  F:=F+1;
PRINT(F,2,0);
  NEWLINE(1);
  TU[3] := 9.25;
  F:=F+1;
PRINT(F,2,0);
  NEWLINE(1);
  TU[4] := 9.25;
  F:=F+1;
PRINT(F,2,0);
  NEWLINE(1);
  TU[5] := -2.5;
  F:=F+1;
PRINT(F,2,0);
  NEWLINE(1);
  HGPLINET(ST,TU,5,+1);
  F:=F+1;
PRINT(F,2,0);
  NEWLINE(1);
  STRARR(UV,LITA,('ANGLES % IN% THE% PHASE% PLANE'));
  F:=F+1;
PRINT(F,2,0);
  NEWLINE(1);
  HGPSYMBLT(-3,0,8,25,0.25,UV,0.0,25);
  F:=F+1;
PRINT(F,2,0);
  NEWLINE(1);
  HGPSYMBLT(2,5,8,75,0.125,FG,0.0,13);
  F:=F+1;
PRINT(F,2,0);
  NEWLINE(1);
  HGPLOTT(0,0,0.0,3,0);
  F:=F+1;
PRINT(F,2,0);
'END';

WRITETEXT('('('('('0')' SUCCESS1')')');
'COMMENT' ANGULAR VELOCITY DISTRIBUTIONS;
'BEGIN';
  HGPLOTT(-10,0,0.0,0,4);
F:=F+1;
NEWLINE(1);
PRINT(F,2,0);
  HGPLINET(YZ,XY,6,+1);
F:=F+1;
NEWLINE(1);
PRINT(F,2,0);
  SYRRP(ABC,LITA,('DEGREES% PER%SEC*%T '));
  HGPAxisVT(-3,0,0.0,ABC,17,6.0,0.0,-720,180,0.75,-1);
  HGPAxisVT(0.0,0.0,RS,11,6.0,90.0,0.0,1,0.6,-1);
  STRARR(DEF,LITA,('ANGULAR % VELOCITY % DISTRIBUTION'));
  HGPSYMBLT(-3,0,8,25,0.25,DEF,0.0,29);
  HGPSYMBLT(2,5,8,75,0.125,FG,0.0,13);
  HGPLINET(ST,TU,5,+1);

```



```

      'COMMENT' AMPLITUDE DISTRIBUTION;
'BEGIN';
      HG PLOTT(+20,0,15.5,0,4);
F:=F+1;
NEWLINE(1);
PRINT(F,2,0);
      HG PLINET(MNO,JKL,6,+1);
F:=F+1;
NEWLINE(1);
PRINT(F,2,0);
      HG PAXISVT(-3,0,0,0,BCD,0,6,0,0,0,-10,1,0.3,-1);
      HG PAXISVT(0,0,0,0,RS,11,6,0,90,0,0,0,1,0.6,-1);
      STRARR(XYZA,LITA,('AMPLITUDE % DISTRIBUTION'));
      HG PSYMBLT(-3,0,8,25,0.25,XYZA,0,0,22);
      HG PSYMBLT(2.5,8,75,0.125,FG,0,0,13);
      STRARR(XYZC,LITA,('AMPLITUDE'));
      HG PSYMBLT(-0.75,-0.6,0.125,XYZC,0,0,9);
      HG PLINET(ST,TU,5,+1);
      HG PLOTT(0,0,0,0,3,0);
'END';
      CLOSEPLOT;
'END';
'END';
'END';

```

```

JOB PHASEPLANE,ASSTF215,WHITEHEAD,STREAM 4  RUN BY GEORGE 2/MKSC
DISPLAY ' 4PHASEPLANE '
LOG
MAXTIME 600
VOLUME 4000
BULLETIN
! PROGRAMLOAD INFO
ALGOLUB 2PHASEPLANE,600,SECS,3000,LINES
! PROGRAMLOAD XALE,COMP
! CONSOLIDATE
! RUN
****

```

ACCOUNT CODE	ASSTF215	DATE	16/07/70
JOB NAME	PHASEPLANE	START TIME	21/11/51
USER NAME	WHITEHEAD,ST	END TIME	21/16/16

PERIPHERALS USED: 32

TOTAL MILL TIME	78
INPUT RECORDS	913
OUTPUT RECORDS	613

PHASE PLANE REPRESENTATION OF A WAVEFORM

BEGIN REAL A,B,C

INTEGER I,J,K,L,M,N,CHARAC

SWITCH S:= RECYCLE,POIS

SETORIGIN(550,6,6,1)

AXES(5,5,10,10,10,10)

PLOTTER(5,1)

FOR I:=-50 STEP 5 UNTIL 50 DO

BEGIN CHARAC:=(I/5)

MOVEPEN(I-2,3)

PRINT DIGITS(2), CHARAC

END

PLOTTER(5,3)

FOR I:=-100 STEP 5 UNTIL 100 DO

BEGIN CHARAC:=(I/5)

MOVEPEN(-3,I)

PRINT DIGITS(2), CHARAC

END

MOVEPEN(-50,120)

PRINT PLOTTER(10,1), I G.D.WHITEHEAD PHASE PLANE PLOT ?

MOVEPEN(-40,6)

PRINT PLOTTER(5,1), I AMPLITUDE ?

MOVEPEN(-6,80)

PRINT PLOTTER(5,3), I SLOPE ?

READ N,A,B

M:=2

C:=(A-B)

RECYCLE : A:=B'

B:=0'

READ B'

C:=A-B'

DRAWLINE(B,C)'

M:= M+ 1'

IF M LESSEQ N THEN GOTO RECYCLE ELSE GOTO POTS'

POTS : PRINT E END ?'

END'

END'

A P P E N D I X C

ESTIMATE UNCERTAINTY IN WAVEFORM ANALYSIS

In the practical analysis of continuous random variables or random processes, it is necessary to know how to control or specify the measurement accuracy of the analysis data that has been generated. If the measurement accuracy is not known the analysis data reduces to qualitative data.

Basically there are two sources of measurement error in random analysis which must be treated separately, for the purposes of error or uncertainty analysis. When the uncertainty analysis is complete and the analysis data collected the two separated uncertainties or errors may be compounded into a single statement of accuracy.

Firstly there is the basic measurement or calibration error. A statement of this error describes the accuracy with which the equivalent electrical voltage or waveform represents the actual quantity measured. Basic measurement and calibration error may be controlled and measured through the proper choice of instrumentation and calibration procedures. However, it is not the purpose of this appendix to discuss instrumentation and calibration, no further comment will therefore be made on this topic. Some details of the basic measurement accuracy of random data analysis is given in Table II of reference 2.

Secondly, and equally important, there is always associated with the analysis of random raw data (whether in the Amp, time or frequency domain) a statistical or estimate uncertainty which is conceptually different to

basic measurement and calibration uncertainty. To understand statistical uncertainty it is essential to think in terms of populations and samples taken from populations.

With deterministic vibration the amplitude time history record will repeat itself exactly if the population of deterministic vibration from which it came is sampled at different times. If the population of random vibration is sampled at different times each sample will be unique and is never likely to be exactly the same as any other sample. Since different samples from the same population can never be exactly the same, it is reasonable to suggest that an average description of each sample would also differ from sample to sample. Further, if the average value of two very long samples were compared we would expect the differences to be less than if we were to compare an average description for the population or universe of discourse, with the average description of a long sample and then a short sample, we would expect to find that the difference was less in the case of the long sample. This expected difference between the population or true values and the sample or estimated values is what is referred to as statistical measurement uncertainty.

These concepts will now be expanded into the theory of statistical estimates. Terms such as bias error, variance error, standard error and confidence limits, which are the language of statistical uncertainty will also be explained. An attempt will also be made to locate the frontier of knowledge in uncertainty analysis based upon current research work.

STATISTICAL ESTIMATION THEORY

If $\{x(\alpha, t)\}$ represents a set of estimates where the range of t and α are $-\infty < t < \infty$, $1 \leq \alpha \leq \infty$. If U is a true average value of the random process $\{x(\alpha, t)\}$ such as the mean value, auto-correlation value or power spectral density value. Suppose that $\hat{U}(\alpha, T)$ is an estimate of U obtained by time averaging a single sample function (for a description of the structure of random processes see reference 1) over a finite time period T . If the ensemble of time average values is $\{\hat{U}(\alpha, T)\}$, if this ensemble is averaged this average value should equal the true value U without error. A set of estimates having this property is said to be unbiased, i.e. a set of estimates $\{\hat{U}(\alpha, T)\}$ is said to be unbiased if independent of T . The expected value is the true value.

i.e.

$$E_{\alpha} \left\{ \hat{U}(\alpha, T) \right\} = U$$

When this happens $\hat{U}(\alpha, T)$ is said to be an unbiased estimate of U .

It should be realised, however, that the mere fact that this ensemble average $E_{\alpha} \left\{ \hat{U}(\alpha, T) \right\}$ for fixed T , is unbiased does not imply that any one of the sample functions $\hat{U}(\alpha, T)$ will have a value close to the true value. Wide deviations from the true value might be observed. To describe these deviations the mean square value is used.

i.e.

$$E_{\alpha} \left\{ \hat{U}(\alpha, T) - U \right\}^2$$

If this mean square error approaches zero as $T \rightarrow \infty$ or becomes large, then any sample function estimate $\hat{U}(\alpha, T)$ would tend to closely approximate its true value U . Any set of estimates having this property are known as

consistent estimates.

i.e. A set of estimates is consistent if

$$\lim_{T \rightarrow \infty} E_{\alpha} \left\{ \hat{u}(\alpha, T) - u \right\}^2 = 0$$

when this happens $\hat{u}(\alpha, T)$ is a consistent estimate.

Digressing for a moment, note that the mean square value, variance value and mean value are related by the following:

Mean Square Value = Variance + Square of
Value Mean Value

$$\text{i.e. } \bar{x}^2 = \sigma_x^2 + [\bar{x}]^2$$

In these terms the mean square error is the sum of variance error and the bias or mean value error squared, thus:

$$E_{\alpha} \left\{ \hat{u}(\alpha, T) - u \right\}^2 = \sigma^2 \left\{ \hat{u}(\alpha, T) \right\} + \left[b \left\{ \hat{u}(\alpha, T) \right\} \right]^2$$

$$\begin{aligned} \text{where } \sigma^2 \left\{ \hat{u}(\alpha, T) \right\} &= E_{\alpha} \left[\hat{u}(\alpha, T) - E_{\alpha} \left\{ \hat{u}(\alpha, T) \right\} \right]^2 \\ &= E_{\alpha} \left\{ \hat{u}^2(\alpha, T) \right\} - \left[E_{\alpha} \left\{ \hat{u}(\alpha, T) \right\} \right]^2 \end{aligned}$$

$$\text{and } b \left\{ \hat{u}(\alpha, T) \right\} = \left[E_{\alpha} \left\{ \hat{u}(\alpha, T) \right\} - u \right]$$

In practical terms it is the job of the Test Engineer and Data Analyst to ensure that this bias and variance error will approach zero as T. Expressions for the bias and variance error in practical analysis of random data are given shortly.

A third property that is desirable of estimates is that they should be efficient. Estimate efficiency is difficult to define mathematically and if defined rarely leads to a clearer understanding of the term. In words, however, an efficient estimate may be thought of as an estimate chosen from a population of many, where all estimates have the same mean or expected value. The most efficient estimate being the one having the least variance and the least efficient estimate being the estimate with the greatest variance.

Some examples of efficient estimates are the arithmetic mean of a sample made up of N observations

i.e. $\bar{x} = \frac{1}{N} \sum_{i=1}^N x_i$ is an efficient estimation of the population mean μ . The sample variance S^2 .

i.e. $S^2 = \frac{1}{N} \sum_{i=1}^N (x_i - \bar{x})^2$

is an efficient estimation of the population variance σ^2 .

Note, however, that whilst the expected value of the sample mean value

$$E[\bar{x}] = \mu$$

indicates that the sample mean value is an unbiased estimate of μ (since by definition the expected value is the true value) the sample variance, S^2 , is a biased estimate of the population variance σ^2 because

$$E[S^2] = \left(\frac{N-1}{N}\right) \sigma^2$$

To remove the bias of the sample variance, it may be corrected as follows:

$$E \left[\left(\frac{N}{N-1} \right) S^2 \right] = E \left[\frac{1}{N-1} \sum_{i=1}^N (x_i - \bar{x})^2 \right] = \sigma^2$$

(A proof of this relationship is given by Weatherburr (1961) on page 131. From this expression, note that the bias in estimating the variance is small for large samples but large for small samples. Correction of Variance bias is therefore usually reserved for small samples (say 30 or below).

MEASUREMENT UNCERTAINTY OF ESTIMATES IN TERMS OF STANDARD ERROR AND CONFIDENCE INTERVAL

The variance in a set of estimates has already been defined as the second moment of the estimate values about the estimate mean value.

i.e.

$$S^2 = E_{\alpha} \left[\hat{u}(\alpha, T) - E_{\alpha} \left\{ \hat{u}(\alpha, T) \right\} \right]^2$$

or for simplicity
$$S^2 = E \left[(x_i - \bar{x}) \right]^2$$

where x_i are the individual estimates and \bar{x} the estimate mean value. It is usual when dealing with samples and populations to denote the sample or estimate variance as S^2 and the population or true value as σ^2 .

Note that for a set of estimates having a zero mean value the variance is identical to the mean square value.

i.e.

$$E \left[x_i^2 \right] = S^2, \quad \text{for } \bar{x} = 0.$$

Again it is usual to call the sample an estimate, mean value \bar{x} and the true value μ .

The standard deviation of a set of estimates is the positive square root of its variance, hence, for a set of estimates:

$$\text{Standard Deviation} = \sqrt{S^2} = S$$

For the population from which the sample came and for which S is an estimate, the standard deviation is σ .

The standard deviation provides only a simple description of the variation in a set of estimates. A better description of variation in a set of estimates is obtained by studying the distribution of estimates. If the distribution of estimate values happens to follow a NORMAL distribution certain conclusions may be drawn from this distribution by studying the standard deviation in the estimates.

If the estimates $\hat{U}(\alpha, \tau)$ of the true value U follow a normal distribution then it can be said that about two thirds of the estimates, $\hat{U}(\alpha, \tau)$ will have values lying within $\pm S$ of U . More precisely we can expect to find U lying within the intervals $\hat{U}(\alpha, \tau) \pm S$, $\hat{U}(\alpha, \tau) \pm 2S$ and $\hat{U}(\alpha, \tau) \pm 3S$ for about 68.27%, 95.45% and 99.73% of the number of estimates respectively. (Note that the number of estimates should be greater than 30 in order to make the assumption of normality reasonably valid.)

In the Analysis of Uncertainty to follow use will be made of a measure called the Standard Error or Normalised Standard Deviation. Standard Error carries exactly the same interpretation as standard deviation, as given above in conjunction with estimate distributions. It is the dimensionless equivalent of standard deviation.

At this point it would be valid to ask what has statistical estimation theory to do with waveform analysis. The answer is nothing if you are ~~not~~ thinking of the waveform as a single time history record and not as a sample from an everlasting population. What must be done to ensure that estimation theory has a place in your thought is to construct a Random Process of the type shown in Figure 4, Reference 1. This may be done by conducting a single experiment and acquiring a long time history record of the variable during this experiment. In this case the random process would be made artificially by dividing the long record into a large number of smaller records each of equal length and placing them side by side. Alternatively a more realistic random process may be constructed by repeating an experiment a large number of times under the same controlled conditions (Reference 1, page 156 explains how a Random Process arises). With a Random Process in mind it should be possible to grasp the importance of estimates, uncertainty and sets, if the Random Process is now thought of as a sample from a population. The population existing not only now but well into the future.

Whilst Random Processes provide a foundation for generating predictive analysis data it must be admitted that a process or ensemble is not always a necessity. In particular in diagnostic analysis and most descriptive analysis sufficiently accurate estimates of the chosen descriptive functions

can be extracted from sufficiently long single samples. Since these later modes of analysis probably form 90% of all analyses carried out in waveform analysis more emphasis will now be placed upon the uncertainty in estimating population values from single samples. It must be clearly understood however, that no statistical statements in the true sense can be made based on a sample size of one.

The preceding notes should provide a reasonable background for what is to follow. If they have not been understood the reader is strongly recommended to read Spiegel (1961), Chapter 9 and Weatherburn (1961), Chapters 3 and 6.

STATISTICAL DEGREE OF FREEDOM

Statistical degree of freedom is defined now because it will be used extensively throughout the following treatments. It represents the equivalent number of events or maximum number of individual events contained in a continuous random variable. It is calculated from the equivalent ideal bandwidth B_n /s of the signal and the length of the record. Since a continuous random variable may be completely reconstructed from digital samples taken at $1/2B_n$ seconds apart the Statistical degree of freedom of a waveform of length T seconds is given by

$$n = 2BT.$$

The equivalent ideal bandwidth B_n and equivalent noise bandwidth of a waveform are defined in Reference 2, page 33.

of τ_w by some distribution not yet known.

The number of statistical degrees of freedom for each measurement of τ_w is given by :

$$n = 2B_N \tau_w$$

Since $\tau_w = WT \hat{\rho}(x)$ from above

$$n = 2WTB_N \hat{\rho}(x)$$

B_N is the equivalent ideal noise bandwidth of the time history record.

The mean square standard error is here defined for the measurement $\hat{\rho}(x)$ as the ratio of the variance in the measurement $\sigma^2 \hat{\rho}(x)$ (for zero mean data) to the square of the measurement value $\hat{\rho}^2(x)$

i.e.

$$\epsilon^2 = \frac{\sigma^2 \hat{\rho}(x)}{\hat{\rho}^2(x)}$$

In terms of the true population variance, $\sigma^2 \hat{\rho}(x)$ the mean square error is given by:

$$\epsilon^2 = \frac{\sigma^2 \hat{\rho}(x)}{n \hat{\rho}^2(x)} \quad \text{since} \quad \sigma^2 \hat{\rho}(x) = \frac{\sigma^2 \rho(x)}{n}$$

Assuming $\sigma^2 \rho(x) \doteq \hat{\rho}^2(x)$ we now obtain

$$\epsilon^2 = \frac{1}{n} = \frac{1}{2WB_N T \hat{p}(x)}$$

or

$$\epsilon = \left(\frac{1}{2WB_N T \hat{p}(x)} \right)^{1/2}$$

This expression gives the most conservative estimate of Standard Error. It was stated that the measurement of t_w will be distributed in some manner about the true value of t_w . This is now equivalent to saying that $\hat{p}(x)$ will be distributed in some manner about $p(x)$. It might be assumed that $\hat{p}(x)$ is normally distributed with a standard error of ϵ . This would only be true if the original random variable was also normally distributed. However since the original random variable was time, i.e. t_w in $\hat{p}(x) = 1/TW \int_0^T t_w dt$ and since time cannot physically take on negative values t_w could not have been normally distributed. To the rescue comes the Central Limit Theorem (see Bendat, 1958, page 96). This states, in essence, that the distribution of $\hat{p}(x)$ will approach normality about $p(x)$ as the number of degrees of freedom becomes large. This is true regardless of the distribution of the original random variable. In terms of the standard error, if ϵ is less than unity, then $\hat{p}(x)$ may be considered to be normally distributed about a mean of $p(x)$ with normalised standard deviation ϵ .

When this occurs specific confidence statements can be associated with ϵ using a normal distribution table. Note that the standard error of the

measurement becomes large as $\hat{p}(x)$ becomes small. This imposes a great restriction upon the generality of the above uncertainty. Realising this, Bolt, Deranek and Newman Inc. conducted an experimental programme to try to establish a more general relationship. Their empirical expression for ϵ will be given shortly.

To determine if the above expression for ϵ gives a valid estimate of a Normal distribution, standard error ϵ should not be allowed to exceed 0.3. This discussion may be clarified by the following example.

Assume we have an analogue amplitude probability analyser and a time history record of say random vibration. Let the length of the record be 10 seconds. Assume also that this sample record has a uniform power spectrum with a sharp cut off at 500 c/s, by a low pass filter. That is, the Equivalent Ideal Noise bandwidth B_N is 500 c/s. Let the analogue analyser have an amplitude window equal to $1/10$ th of the rms voltage of the sample record. If the rms voltage is one volt, then the window is 0.1 volts. Note that the rms voltage is also the standard deviation of the record, which is thus unity. If the centre of the amplitude window is now set to some voltage x_1 and the time t_1 is measured that the signal spends in the window, if this is divided by T and W an estimate $\hat{p}(x)$ is obtained. If the centre voltage is now moved 0.1 volts a second probability density estimate will be obtained. The procedure is repeated until the amplitude range of interest has been covered: say $\pm 4 \sigma = \pm 4$ volts.

From Table V of Reference 2 the minimum analysis time is given as

$t = T_a \times A/W = 10.8/0.1 = 800$ seconds. The amplitude range could be

swept or stepped. If swept the scan rate should be no greater than

$W/T_a = 0.1/10 = 0.01$ volts per second.

The standard error of the estimates will be given by:

$$\epsilon = \left[\frac{1}{2(500)(10)(0.1)\hat{p}(x)} \right]^{1/2}$$

$$\therefore \epsilon = \left(\frac{0.001}{\hat{p}(x)} \right)^{1/2}$$

i.e. the standard error at any centre voltage is a function of the probability density measured at that voltage. For the purpose of the example and simplicity, assume that the measured probability density function followed a normal distribution. Suppose that the measurements at principle points on the normal distribution and their ϵ values were as follows. (See Figure)

Centre voltage of window	Measured value for $p(x)$	
0	0.4	0.05
+1	0.24	0.064
+2	0.054	0.14
+3	0.0044	0.48
+4	0.0001	3.2

Consider a 68% confidence interval, i.e.

$$P_{\text{rob}} \left[p(x) - \epsilon \hat{p}(x) \leq \hat{p}(x) < p(x) + \epsilon \hat{p}(x) \right] = 0.68$$

In words the probability that the estimated value $\hat{p}(x)$ will fall within the range of $\pm \epsilon$ of the true value is 0.68.

The 68% confidence interval for the true probability density at zero volts is the range of $\hat{p}(x) \pm \epsilon$ between $0.4 - 5\%$ of 0.4 and $0.4 + 5\%$ of 0.4 , i.e. from 0.38 to 0.42 . At one volt we can be 68% confident that $p(x)$ is between 0.22 and 0.28 . At two volts for a 68% confidence interval $p(x)$ is between 0.046 and 0.062 . At three or four volts the error has become too large (greater than 0.3) to allow an assumption that the distribution of $p(x)$ is normal. At four volts ϵ indicates a 68% confidence range of $1 - 0.0002$ to $+ 0.0004$ for $p(x)$. Since $p(x)$ cannot be negative this latter statement has no meaning. Therefore at three and four volts the statistical accuracy is not known. The three and four volt centre voltages in the example provide an opportunity for specifying a statistical accuracy. Since the existing 10 seconds of raw data have yielded statistical uncertainty so wide as to be indefinable we will reverse the procedure and find the length of record based upon an acceptable uncertainty (say $0.3 = \epsilon$).

At three volts $\hat{p}(x)$ is 0.0044 . Therefore if ϵ is established as 0.3

$$0.3 = \left[\frac{1}{2(500)(0.1)(0.0044)(T)} \right]^{1/2}$$

solving for T yields : $T = 25.2$ seconds. At four volts $\hat{p}(x)$ is 0.0001

$$\therefore 0.3 = \left[\frac{1}{2(500)(0.1)(0.0001)(T)} \right]^{1/2}$$

and T is 1110 seconds = 18.5 mins.

This concludes the example.

Bolt Beranek and Newman (1960) have derived an expression for standard error based upon threshold crossing theory. Their expression (for a guide to the derivation see Rice, 1944) given below, makes the assumption that the random waveform has a normal amplitude probability density function.

$$\epsilon = \left[\frac{1}{2.89 B_N T W \hat{\rho}(x)} \right]^{1/2}$$

When this relationship was checked experimentally using a random noise generator with a sharp frequency cut off, it was found that the results led to the empirical relationship.

$$\epsilon = \left[\frac{1}{\pi^3 B_N T W \hat{\rho}(x)} \right]^{1/2}$$

Thus the experimental derivation indicated that the standard error was about half that predicted by their theoretical formulae and less than half that predicted by the previously derived expression.

Note that each expression is identical except for the coefficients which were 0.707, 0.588 and 0.254.

Bendat and Piersol (1965) have suggested a more general relationship for which is based upon the spectral characteristics of the data and upon the way in which the data was analysed. Their general relationship is given by:

$$\epsilon = \left[\frac{C_0}{X T W \hat{\rho}(x)} \right]^{1/2}$$

where $X = B_N$ or \bar{v}_0

and C_0 is a coefficient dependent upon the frequency composition and way in which the estimate $\hat{p}(x)$ is obtained.

\bar{v}_0 is the expected number of zero crossings per second.

In the previous expressions for ϵ the frequency bandwidth of the signal was assumed to range from zero to some upper frequency f_3 the bandwidth from zero to f_3 being the equivalent noise bandwidth B_N c/s. Bendat and Piersol have found experimentally that for signals having a relatively wide bandwidth, X , in the above expression for ϵ , is \bar{v}_0 . They have evaluated the coefficient C_0 for this case as 0.26, i.e.

$$\epsilon^2 = \frac{0.068}{\bar{v}_0 T W \hat{p}(x)}$$

$$\text{or } \epsilon = \left[\frac{0.068}{\bar{v}_0 T W \hat{p}(x)} \right]^{1/2}$$

For the case of signals with relatively narrow bandwidths $X = B_N$ in the above expression. The coefficient for this case was found to be 0.17, i.e.

$$\epsilon^2 = \frac{0.028}{B_N T W \hat{p}(x)}$$

$$\text{or } \epsilon = \frac{0.17}{(B_N T W \hat{p}(x))^{1/2}}$$

It is thought that the first expression applies when the ratio B_N / \bar{v}_0 is greater than a third. This applies to low frequency signals where the bandwidth is narrow but the expected number of zero crossings is small. The second expression with a coefficient of 0.17 is thought to apply to signals having a ratio which is smaller than a third. This happens when a narrow bandwidth signal is centred on a high frequency.

In human factors research it is unlikely that frequencies will be encountered much above 1000 c/s. The range of frequency of interest is thus 0-1000 c/s. Since 1000 c/s may be regarded as a fairly low frequency it can be seen that in human factors research, waveforms most closely approximate to the range of frequency specified by the expression

$$\epsilon = \frac{0.26}{(\bar{v}_0 T W \hat{p}(x))^{1/2}}$$

The value of \bar{v}_0 for signals having a uniform power spectrum over the bandwidth B_N c/s which extends to zero frequency is given by,

$$\bar{v}_0 = 1.15 B_N$$

If the power spectrum has a sharp peak at some centre frequency f_c then,

$$\bar{v}_0 = 2f_c$$

If the power spectrum is uniform between f_a and f_b

$$\bar{v}_0 = \frac{2}{\sqrt{3}} \left\{ f_a^2 + f_a f_b + f_b^2 \right\}^{1/2}$$

The generally expected number of zero crossings for any signal may be deduced from the power spectrum of the signal thus,

$$\bar{v}_0 = 2 \left\{ \frac{\int_0^\infty f^2 G_x(f) df}{\int_0^\infty G_x(f) df} \right\}^{1/2}$$

For most physical analyses of amplitude probability density, if the uncertainty criterion is $\epsilon = 0.1$ or greater, then it will be reasonably accurate to use a coefficient value somewhere between 0.26 and 0.17,

say 0.2. Bandwidth rather than expected zero crossings may then be used as the frequency description. It is therefore suggested that for most human factors research the expression:-

$$\epsilon = \frac{0.2}{(B_N T W \hat{p}(x))^{1/2}}$$

will be adequate. A plot of ϵ with $C_0 = 0.2$ is given in Figure versus the product $B_N T W$ for various values of $\hat{p}(x)$.

By now it will be apparent that some a priori knowledge of the waveform noise bandwidth and the amplitude probability density values of the waveform are needed to specify a value for ϵ . Since this is not possible, for the most part, it is thus essential to pre-analyse the waveform in the amplitude domain to ascertain the set of values $\hat{p}(x)$. Additionally the equivalent noise bandwidth must also be known.

MEAN SQUARE UNCERTAINTY

So far, what has been considered is only the variance term (normalised) of the mean square estimate error. This was mentioned during discussion on statistical estimation theory. The mean square error is given by:

$$\begin{aligned} \text{M.S. Error} &= \text{Variance} + (\text{Bias})^2 \\ &\text{Error} \quad \quad \quad (\text{Error}) \\ &= \epsilon^2 + (\text{Bias})^2 \end{aligned}$$

The bias error may be approximated by:

$$b = \frac{W^2}{2} \hat{p}''(x)$$

$\hat{p}''(x)$ being the second derivative of the amplitude probability density

estimate (the derivation of the bias error in amplitude estimates is similar to that for power spectral density estimates, a derivation for which will be given later).

Thus the total mean square error is given by:

$$\text{M.S. Error} = \frac{0.04}{B_N T W \hat{\rho}(x)} + \left[\frac{W^2}{24} \hat{\rho}^{(00)}(x) \right]^2$$

$$\text{i.e. } E \left[\hat{\rho}(x) - \rho(x) \right]^2 = \frac{0.04}{B_N T W \hat{\rho}(x)} + \frac{W^4}{576} \left[\hat{\rho}^{(00)}(x) \right]^2$$

CONCLUSIONS REGARDING MEASUREMENT UNCERTAINTY IN AMPLITUDE PROBABILITY DENSITY ESTIMATES

A general expression for the variance term given as the standard error is not yet either theoretically established or empirically suggested.

The general expression will be a complicated function of frequency bandwidth and expected number of zero crossings per second.

The best approximates, theoretically derived and experimentally substantiated is that given by Bendat and Prensels as

$$E = 0.2 \quad \epsilon = \frac{0.2}{B_N T W \hat{\rho}(x)}$$

A suggested criterion for ϵ based upon values used by NASA is $\epsilon = 0.01$. This gives 68% confidence that the true value $p(x)$ lies within $\pm 1\%$ of the measured value.

To specify the sample length of record needed to maintain the measurement uncertainty ϵ for a given amplitude window and frequency bandwidth B_N c/s it is necessary to know the value of the measured amplitude probability density $\hat{p}(x)$. Hence, since $\hat{p}(x)$ is an unknown during the experimental planning stage, the required record length cannot be predicted without carrying out some analysis. The mutual interaction of ϵ and T can therefore be removed only by a process of iterative measurement and analysis. The contribution made to the mean square error by the bias error will be very small if W is small. In comparison with the variance error it will be negligible and may therefore be neglected. In general terms the bias error may be thought of as that due to an amplitude window of finite value and the variance as that due to a sample record of finite length.

MEASUREMENT UNCERTAINTY IN SECOND ORDER OR JOINT AMPLITUDE PROBABILITY DENSITY ESTIMATES

No published experimental work on the measurement uncertainty in joint APD estimates is known to the author.

Example: Physical values of the 1st order amplitude probability density function.

Often it helps to know the range of physical values that a measuring instrument should read for a given input. This is basically the

calibration problem but some numerical values often help fix some of the ideas as well. The example chosen is an analysis of the amplitude probability density of the output from a random noise generator. The frequency bandwidth of the random noise generator was 100 c/s to 8000 c/s the r.m.s. level of the output was 1.00 volt and the amplitude window was 0.100 volt. Figure shows the amplitude probability density for the random noise plotted against voltage amplitude (also in this case, standard deviation).

The analyser was a Bruel and Kjaen Model 160 and the noise generator a General Radio Co. Model 1390A. Fitted to this measured curve is a true normal distribution. The measured distribution is shown to have a very slight positive skew but is otherwise truly normal. The physical values of $\hat{p}(x)$ for this data have a maximum of 0.4 and tails off to zero at about $3\frac{1}{2}\sigma$. The measurement uncertainty at any point on the curve is given by:

$$\epsilon = 0.2 / (W B_n T \hat{p}(x))^{1/2}$$

where $W = 0.1$ volt

$T = 4.6$ seconds = Averaging Time

$B_n = 7900$ c/s

$$\epsilon = \frac{0.0042}{\hat{p}(x)}$$

For example the interval in which we are 68% confident that the true value will lie at a value $\hat{p}(x) = 0.4$ is

$$\epsilon = \frac{0.0042}{0.4} = 0.00665$$

i.e. the true value at $\hat{p}(x) = 0.4$ lies in the range $\hat{p}(x) = 0.39734$ to 0.40665 . At $\hat{p}(x) = 0.05$

$$\epsilon = 0.01875$$

\therefore the range in which the true value will lie with 68% confidence is $\hat{p}(x) = 0.04999064$ to 0.05000936 . This concludes the example.

MEASUREMENT UNCERTAINTY IN FREQUENCY ANALYSIS

Power spectral density - weakly stationary

An estimate of the true mean square spectral density function $G_x(f)$ is given by:

$$\hat{G}_x(f) = \frac{1}{BT} \int_0^T x_B^2(t) dt$$

where $x_B^2(t)$ is the square of the amplitude time history record passed by the narrow bandpass filter of B c/s.

The mean square measurement error of such an estimate is compounded from a bias error and variance error thus.

$$\text{Mean Square Error} = \text{Variance Error} + (\text{Bias Error})^2$$

i.e.

$$E \left[\hat{G}_x(f) - G_x(f) \right]^2 = E \left[\hat{G}_x(f) - E \hat{G}_x(f) \right]^2 + \left[E(\hat{G}_x(f)) - G_x(f) \right]^2$$

BIAS ERROR

It can be shown that $\hat{G}_x(f)$ provides an asymptotically unbiased estimate of $G_x(f)$ as $T \rightarrow \infty$ provided that as $T \rightarrow \infty$ the bandwidth $B \rightarrow 0$.

This condition on B as a function of T is usually assumed, although of course, not true. By definition the bias error is given by:

$$b \left[\hat{G}_x(f) \right] = E \left[\hat{G}_x(f) \right] - G_x(f)$$

At this point it is necessary to indulge in deep mathematics to obtain the following result. The result for the bias analysis is expressed in the important asymptotic formula which assumes that the autocorrelation function $R_x(\tau)$ of the signal, $\tau R_x(\tau)$ and $\tau^2 R_x(\tau)$ are absolutely integrable over the range $-\infty < \tau < \infty$.

Thus
$$\lim_{T \rightarrow \infty} B^{-2} \left| b \left[\hat{G}_x(f) \right] \right| \doteq \frac{1}{24} \left| G_x^{(2)}(f) \right|$$

$G_x^{(2)}(f)$ is the second derivative of $G_x(f)$ with respect to f . In terms of the equivalent autocorrelation function $G_x^{(2)}(f)$ is given by Bendat (1958) as

$$G_x^{(2)}(f) = -4 \pi^2 \int_0^{\infty} \tau^2 R_x(\tau) e^{-j 2 \pi f \tau} d\tau$$

From the above term for bias it follows that the bias error approaches zero as $B \rightarrow 0$ and the record length $T \rightarrow \infty$.

Making the assumption that T is large leads not unreasonably to the final result:

$$|b[\hat{G}_x(f)]| \doteq \frac{B^2}{24} |G_x(f)|$$

VARIANCE ERROR

$$\sigma^2[\hat{G}_x(f)] = E[\hat{G}_x(f) - E(\hat{G}_x(f))]^2$$

After a lengthy derivation it has been shown that as

$$\begin{aligned} \lim_{T \rightarrow \infty} BT \sigma^2[\hat{G}_x(f)] &\doteq G_x^2(f), \text{ for } f \neq 0 \\ &\doteq 2G_x^2(f), \text{ for } f = 0 \end{aligned}$$

If T is large

$$\begin{aligned} \sigma^2[\hat{G}_x(f)] &\doteq \frac{G_x^2(f)}{BT}, \text{ for } f \neq 0 \\ &\doteq \frac{2G_x^2(f)}{BT}, \text{ for } f = 0 \end{aligned}$$

This equation indicates that the variance approaches zero as $T \rightarrow \infty$

provided that $BT \rightarrow \infty$.

MEAN SQUARE ERROR

$$E[\hat{G}_x(f) - G_x(f)]^2 = \sigma^2 [\hat{G}_x(f)] + [b(\hat{G}_x(f))]^2$$

$$= \frac{\hat{G}_x^2(f)}{BT} + \left[\frac{B^2}{24} |G_x(f)|^2 \right]^2$$

From the variance the standard error may be defined as the ratio of the variance in the estimate to the square of the true value.

Hence

$$E^2 = \frac{\sigma^2 [\hat{G}_x(f)]}{G_x^2(f)} = \frac{1}{BT} \frac{G_x^2(f)}{G_x^2(f)}$$

$$\therefore E^2 = \frac{1}{BT}$$

An alternative derivation of the variance term comes from the statistical degree of freedom of the signal for which $\hat{G}_x(f)$ is an estimate.

The equivalent number of events is contained in a continuous white noise signal of length T secs and bandwidth B c/s is given by

$$n = 2 BT$$

If the signal $x(t)$ is assumed to have a normal amplitude probability density function, the sampling distribution for an estimate $\hat{G}_x(f)$ will be:

$$\hat{G}_x(f) \sim \frac{G_x(f) \chi^2_{n-1}}{(n-1)}$$

\sim means distributed as and χ^2 is a chi-square distribution with $(n - 1)$ degrees of freedom.

The variance in $\hat{G}_x(f)$ is given by

$$\begin{aligned} \text{Variance } \hat{G}_x(f) &= \text{Var} \left[\frac{G_x(f) \chi^2}{n-1} \right] \\ &= \left[\frac{G_x(f)}{n-1} \right]^2 \text{Var } \chi^2 \end{aligned}$$

The variance of a χ^2 distribution is equal to twice the degree of freedom hence:

$$\text{Variance } \chi^2 = 2(n-1)$$

$$\therefore \text{Variance } \hat{G}_x(f) = \frac{2 G_x^2(f)(n-1)}{(n-1)^2} = \frac{2 G_x^2(f)}{(n-1)}$$

substituting $n = 2 \text{ BT}$ into the variance expression derives

$$\text{Variance } \left[\hat{G}_x(f) \right] = \frac{2 G_x^2(f)}{2\text{BT} - 1}$$

with

$$\epsilon^2 = \frac{\text{Var.} [\hat{Q}_x(f)]}{Q_x^2(f)}$$

$$\epsilon^2 = \frac{1}{BT-1} \doteq \frac{1}{BT} \quad \text{for } BT \gg 1$$

This concludes the derivation of uncertainty for the power spectral density. In the expression for mean square error the bias term will be insignificant except for cases where the derivative on slope of the power spectrum is changing rapidly with the bandwidth B. For most cases the measurement uncertainty may be taken as the variance term. Figure shows ϵ plotted against the BT product.

MEASUREMENT UNCERTAINTY OF JOINT POWER SPECTRAL DENSITY FUNCTION

An estimate of the joint or cross power spectral density function for a pair of records is given by

$$\hat{Q}_{xy}(f) = \hat{C}_{xy}(f) - j \hat{Q}_{xy}(f)$$

$$\hat{C}_{xy}(f) = \frac{1}{BT} \int_0^T x_B(t) y_B(t) dt$$

$$\hat{Q}_{xy}(f) = \frac{1}{BT} \int_0^T \overset{\text{wavy}}{x_B(t)} y_B(t) dt$$

Physically, the cross spectral density function must be measured in two parts. The first part given by $\hat{C}_{xy}(f)$ is the Co-Spectrum. The second part given by $\hat{Q}_{xy}(f)$ is the Quadrature Spectrum. To form the cross spectrum the two components are subtracted "out of phase". To compute the

Co and Quad spectrum of $\hat{G}_{xy}(f)$ the signals $x(t)$ and $y(t)$ must both be weakly stationary.

The measurement uncertainty for $\hat{G}_{xy}(f)$ is obtained in the same way as the measure of $\hat{G}_{xy}(f)$. That is by subtracting the measurement uncertainty for the components $\hat{C}_{xy}(f)$ and $\hat{Q}_{xy}(f)$ "out of phase". The measurement uncertainty for the components $\hat{C}_{xy}(f)$ and $\hat{Q}_{xy}(f)$ is the same as for the singular power spectral density function except for a multiplication between $x_g(t)$ and $y_g(t)$. This is equivalent to squaring of $x_g(t)$ in $\hat{G}_x(f)$.

$\overset{w}{x}_g(t)$ in $\hat{Q}_{xy}(f)$ indicates that $x_g(t)$ has been phase lagged 90° relative to $y_g(t)$.

In terms of mean square error, the bias and variance errors for the components are given by:

Bias Error

$$b [\hat{C}_{xy}(f)] \leq \frac{B^2}{24} \left| \hat{G}_{xy}(f) \right|$$

$$b [\hat{Q}_{xy}(f)] \leq \frac{B^2}{24} \left| \hat{G}_{xy}(f) \right|$$

As for the ordinary singular power spectral density function:

$$\left| \hat{G}_{xy}(f) \right| = -4\pi^2 \int_0^{\infty} \tau^2 R_{xy}(\tau) e^{-2\pi f\tau} d\tau$$

Variance Error

Variance $\left[\hat{C}_{xy}(f) \right] \approx \frac{\hat{C}_x(f) \hat{C}_y(f)}{BT}$

Variance $\left[\hat{\Phi}_{xy}(f) \right] \approx \frac{\hat{C}_x(f) \hat{C}_y(f)}{BT}$

In terms of standard error, which is defined for the cross spectral components as

$$E^2 = \frac{\sigma^2 \hat{C}_{xy}(f)}{\hat{C}_x(f) \hat{C}_y(f)} \doteq \frac{1}{BT}$$

and

$$E^2 = \frac{\sigma^2 \hat{\Phi}_{xy}(f)}{\hat{C}_x(f) \hat{C}_y(f)} \doteq \frac{1}{BT}$$

Mean Square Error

$$E \left[\hat{C}_{xy}(f) - C_{xy}(f) \right]^2 = E \left[\left(\hat{C}_{xy}(f) - j \hat{\Phi}_{xy}(f) \right) - \left(C_{xy}(f) - j \Phi_{xy}(f) \right) \right]^2$$

In this form the mean square error in a cross spectral density estimate is very difficult to evaluate. The treatment is beyond the scope of this note.

SPECTRAL BANDWIDTH FOR ORDINARY POWER SPECTRUM

The mean square error for the ordinary singular power spectrum was found to be:

$$E \left[\hat{G}_x(f) - G_x(f) \right]^2 = \frac{G_x^2(f)}{BT} + \left[\frac{B^2}{24} \left| \frac{G_x^{(00)}(f)}{G_x(f)} \right| \right]^2$$

If this expression is now normalised, as before to the mean square standard error ϵ^2 , ϵ^2 is given by

$$\epsilon^2 \left[\hat{G}_x(f) \right] = \frac{E \left[\hat{G}_x(f) - G_x(f) \right]^2}{G_x^2(f)}$$

$$\epsilon^2 = \frac{1}{BT} + \frac{B^4}{576} \left| \frac{G_x^{(00)}(f)}{G_x(f)} \right|^2$$

If $\lambda(f)$ is defined such that,

$$\lambda(f) = \left| \frac{G_x^{(00)}(f)}{G_x(f)} \right|^{1/2}$$

then:

$$\epsilon^2 \doteq \frac{1}{BT} + \frac{B^4}{576} \left[\frac{1}{\lambda(f)} \right]^4$$

$\lambda(f)$ is known as the spectral bandwidth of the signal $x(t)$. $\lambda(f)$ has units of frequency, i.e. c/s ϵ^2 is usually written

$$\epsilon^2 \doteq \frac{1}{BT} + \frac{1}{576} \left[\frac{B}{\lambda(f)} \right]^4$$

In physical terms the spectral bandwidth may be regarded as the frequency bandwidth of the narrowest peak in the power spectral density curve $\hat{G}_x(f)$. $\lambda(f)$ is closely related to the resolution needs in power spectral density analysis. By resolution is meant the ability of an analysis to properly separate adjacent peaks in the power spectrum. If these spectral peaks are one spectral bandwidth apart, then it seems reasonable to select a discriminating filter whose bandwidth B c/s is one half of the spectral bandwidth, i.e.

$$B \leq \frac{\lambda(f)}{2}$$

If B is chosen in this way, the analysis of the data will ensure, with a low probability of error, that if $G_x(f)$ has two peaks $\lambda(f)$ apart that these two peaks can be resolved by taking measurements of $G_x(f)$ at intervals B c/s apart.

Although the spectral bandwidth is really a matter for resolution to be covered in Analysis Specification, it has been introduced here because it will influence the choice of filter bandwidth. Since filter bandwidth for

a given record length will control the measurement uncertainty in the analysis, the presence of $\lambda(f)$ here is valid. The following example may help clarify these ideas.

Assume we have a sample record of random vibration which we wish to describe in the frequency domain. If this record is weakly stationary the power spectral density function $G_x(f)$ may be employed. Suppose that it is desired to estimate $G_x(f)$ with a standard error not exceeding the criterion $\epsilon = 0.01$. That is we want to be 68% confident that the true value $G_x(f)$ for the population $\{x(\alpha, t)\}$ from which the sample record came, lies within $\pm 1\%$ of the estimated value.

To control ϵ we have only to treat B , T and $\lambda(f)$ since the total mean square error (ϵ^2) was given by:

$$\epsilon^2 = \frac{1}{3T} + \frac{1}{576} \left[\frac{B}{\lambda(f)} \right]^4$$

However, whilst B and T may be specified to control ϵ , $\lambda(f)$ is an unknown at the raw data stage. In order to gain some knowledge of $\lambda(f)$ the procedure is to diagnose the raw data by a pilot analysis. To conduct a pilot analysis it is helpful to have some a priori knowledge of the raw data sample. If for instance the raw vibration sample was representative of the whole body vibration of an automobile, we could make a shrewd guess that its power spectrum would have at least two peaks. One peak would occur round about the natural frequency of the body on its road springs and a second around the natural frequency of the unsprung mass on the tyre springs. Suppose the former peak was of order 5 c/s in width, centred around 2 c/s and the latter 4 c/s centred on 12 c/s. We would

chose for the pilot analysis a filter bandwidth of 1 c/s.

The pilot analysis might have revealed a peak in the power spectrum of width 2 c/s. In this case we would maintain the pilot bandwidth of 1 c/s as given by

$$B \leq \frac{\lambda(f)}{2}$$

When B satisfied $B \leq \lambda(f)/2$, the second term in ϵ^2 becomes negligible and ϵ^2 reduces to $1/BT$. Therefore to maintain an E of 0.01 a raw data sample must be acquired of length

$$T = 1/B\epsilon^2 = \frac{1}{1.(0.01)^2}$$

$$T = 10^4 \text{ seconds}$$

$$T = 166.6 \text{ minutes}$$

To illustrate the importance of the relationship $B \leq \lambda(f)/2$ consider the case where the filter bandwidth in the above example was 4 c/s instead of 1 c/s, i.e. $B = 2 \lambda(f)$. If the same record length is used namely 10^4 seconds the mean square standard error becomes:

$$\epsilon^2 = \frac{1}{BT} + \frac{1}{576} \left[\frac{B}{\lambda(f)} \right]^4$$

$$= \frac{1}{4 \cdot 10^4} + \frac{1}{576} \left[2 \right]^4$$

$$= 0.25 \cdot 10^{-4} + 16/576$$

$$= 25 \cdot 10^{-6} + 0.0277$$

Taking the positive square root:

$$\epsilon = 0.005 + 0.167$$

$$\epsilon = 0.172 = 17.2\%$$

Notice how the bias error as given by the second term has swamped the variance error. The measurement uncertainty has been increased from 1% to 17.2%. It is clear from this example how important it is to maintain

$$\beta = \lambda(f)/2 \quad \text{if this is possible. This concludes the example.}$$

CONFIDENCE INTERVAL

The measurement uncertainty in a power spectral density estimate, may alternatively be expressed by a confidence interval. The confidence interval gives the range over which the true value, $G_x(f)$ is known to lie with a given level of confidence. If it is assumed that the distribution of estimated values about the true value follows Chi Square (χ^2) (This also assumes that the original raw data $x(t)$ followed a normal distribution). Then the ratio of the estimated value $\hat{G}_x(f)$, to the true value $G_x(f)$, for a given confidence interval will vary with the degree of freedom of the analysis. Figure illustrates this description of measurement uncertainty. For example if the sample time is 10 seconds and the filter bandwidth

3.5 c/s the degree of freedom is 70 from

$$n = 2BT$$

The 90% confidence limits for a power spectral density estimate with a degree of freedom of 70 is given by:

$$Prab \left[\hat{G}_x(f) - 0.75 \hat{G}_x(f) \leq G_x(f) < \hat{G}_x(f) + 1.37 \hat{G}_x(f) \right] = 0.9$$

From Figure If $\hat{G}_x(f)$ is 2.0 volts²/ c/s then $G_x(f)$ lies in the range 1.5 to 2.73 volts²/ c/s with 90% confidence.

MEASUREMENT UNCERTAINTY IN FREQUENCY RESPONSE ESTIMATES

For a constant parameter linear system having a weakly stationary excitation $x(t)$ and response $y(t)$ and estimate of the frequency response of the system is given by:

$$\hat{H}(f) = \frac{G_{xy}(f)}{\hat{G}_x(f)}$$

$\hat{G}_{xy}(f)$ is the estimated cross power spectral density function between the input and output and $\hat{G}_x(f)$ is the estimated power spectral density function of the input. $\hat{H}(f)$ is the estimated frequency response function which carries a gain factor $|\hat{H}(f)|$ and phase factor $\hat{\phi}(f)$ thus:

$$\hat{H}(f) = |\hat{H}(f)| e^{j\hat{\phi}(f)}$$

The measurement uncertainty in estimating $\hat{H}(f)$ has been studied by Goodman

(1957). Goodman has derived an expression for the measurement uncertainty in the gain and phase factors of $\hat{H}(F)$ based upon the Coherence function $\gamma_{xy}^2(F)$.

where

$$\gamma_{xy}^2(F) = \frac{|G_{xy}(F)|^2}{G_x(F)G_y(F)}$$

Goodman's derivation is given by:

$$P \left[\left| \frac{\hat{H}(F) - H(F)}{H(F)} \right| < \sin \epsilon \text{ and } \left| \hat{\phi}(F) - \phi(F) \right| < \epsilon \right] \\ \doteq 1 - \left[\frac{1 - \gamma_{xy}^2(F)}{1 - \gamma_{xy}^2(F) \cos^2 \epsilon} \right]^n$$

n is the degree of freedom in the measurement given by

$$n = 2BT \quad (B \text{ is the filter bandwidth}).$$

and ϵ is the standard error in estimating the gain or phase factors.

Goodman's equation is shown plotted in Figure . The three sets of curves are plotted for $P = 0.9, 0.85$ and 0.8 with $\epsilon = 0.005, 0.10$ and 0.15 . P and ϵ give the probability of maintaining a desired measurement accuracy for a given degree of freedom in the determination of $\gamma_{xy}^2(F)$. For small ϵ , $\sin \epsilon = \epsilon$, therefore the plotted curves are valid for both the gain and phase factors. For the gain factor the curves represent a measurement uncertainty of 5, 10 and 15%. For the phase factor they represent a measurement uncertainty of 0.05, 0.1 and 0.15 radians.

If it is desired to measure the frequency response function of a linear system with a known accuracy, iterative measures must be used. All that was said about spectral bandwidth etc. when dealing with the ordinary power spectrum applies to the frequency response function. This is because

the measurement of $\hat{H}(f)$ involves $\hat{C}_{yx}(f)$.

To iterate a frequency response function to a given level of measurement accuracy first measure the coherence function for a chosen length of data sample. Next enter the ordinate of Figure at the measured value of $\hat{C}_{yx}(f)$ and follow up to the chosen combination of P and ϵ . Then read off the number of degrees of freedom and calculate a new record length for a given bandwidth. From the new length of data sample estimate the coherence function a second time repeating the procedures until the desired measurement accuracy is realised.

CONCLUSIONS REGARDING THE MEASUREMENT UNCERTAINTY IN FREQUENCY ANALYSIS

For ordinary power spectral density analysis the mean square standard error is given by:

$$\epsilon^2 = \frac{1}{BT} + \frac{1}{576} \left[\frac{B}{\lambda(f)} \right]^4$$

$\lambda(f)$ is the spectral bandwidth of the power spectrum on the bandwidth of the narrowest peak than can be resolved. If possible $\lambda(f)$ should be determined prior to settling the analysis data, i.e. $\hat{C}_{yx}(f)$. It is suggested that $\lambda(f)$ be determined by an iterative procedure. When is known, if possible, the filter bandwidth should maintain the relationship.

$$B \leq \frac{\lambda(f)}{2}$$

Ideally (see Table V, Ref. 3) B should be chosen such that $B \leq \lambda(f)/4$. When B maintains this relationship with $\lambda(f)$ the mean square standard error may be taken as the variance term and the bias terms neglected. If B does

not maintain the above relationship with $\lambda(f)$ the bias error must be included.

The mean square error in estimating the cross power spectral density function is known only for the components $\hat{C}_{xy}(f)$ and $\hat{Q}_{xy}(f)$. The mean square standard error for the components is similar to that for the ordinary power spectral density function, namely

$$\epsilon^2 \left[\hat{C}_{xy}(f) \text{ and } \hat{Q}_{xy}(f) \right] \doteq \frac{1}{BT}$$

For a frequency response function estimate the measurement uncertainty is given by Goodman in terms of phase factor, gain factor, degree of freedom and coherence. Since the frequency response function contains both the ordinary power spectrum and cross spectrum functions all remarks relating to the ordinary and cross spectrums apply to the frequency response function. In order to measure the frequency response function with a known accuracy the coherence function must be determined iteratively.

MEASUREMENT UNCERTAINTY IN TIME ANALYSIS

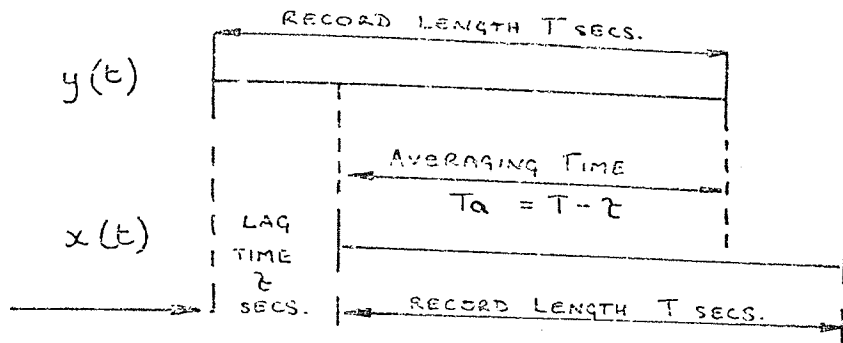
Weakly Stationary - Cross Correlation Function

An estimate of the cross correlation function $R_{xy}(\tau)$ is given by:

$$\hat{R}_{xy}(\tau) = \begin{cases} \frac{1}{T-\tau} \int_0^{T-\tau} x(t)y(t+\tau) dt, & 0 \leq \tau \leq T. \\ \frac{1}{T+|\tau|} \int_{-\tau}^T x(t)y(t+\tau) dt, & -T \leq \tau \leq 0. \end{cases}$$

Note that the integral, for positive τ extends from 0 to $T - \tau$ i.e. the integration time is $T - \tau$ and not T which is the record or sample length of vibrations. The diagram below illustrates this point.

Figure



For most practical purposes, if T is long compared to τ then a "second class" estimate of $R_{xy}(\tau)$ may be used as given by

$$\hat{R}_{xy}(\tau) \doteq \frac{1}{T} \int_0^T x(t) y(t + \tau) dt$$

note that $\hat{R}_{xy}(\tau) \neq \hat{R}_{yx}(-\tau)$. From now on only a positive lag time will be considered. The measurement uncertainty being the same for both positive and negative lag.

MEAN SQUARE MEASUREMENT UNCERTAINTY

The mean square error in an estimate $\hat{R}_{xy}(\tau)$ is given by

$$E \left[\hat{R}_{xy}(\tau) - R_{xy}(\tau) \right]^2 = E \left[\tilde{R}_{xy}(\tau) - E(\hat{R}_{xy}(\tau)) \right]^2 + \left[E(\hat{R}_{xy}(\tau)) - R_{xy}(\tau) \right]^2$$

BIAS ERROR

The second term in the expression for mean square error is the square of the bias error. For a cross correlation estimate the bias error is zero. This is proved below.

$$\begin{aligned}
 E \left[\hat{R}_{xy}(\tau) \right] &= E \left[\frac{1}{T} \int_0^T x(t)y(t+\tau) dt \right] \\
 &= \frac{1}{T} \int_0^T E \left[x(t)y(t+\tau) \right] dt \\
 &= \frac{1}{T} \int_0^T R_{xy}(\tau) dt
 \end{aligned}$$

Since the expected value of an expectancy is the expected value itself (the terms, integral on average, may replace the expectancy term in continuous data)

$$E \left[\hat{R}_{xy}(\tau) \right] = R_{xy}(\tau)$$

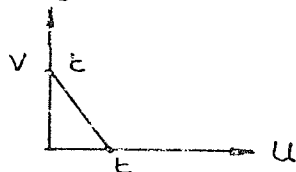
Therefore, by definition if the expected value of an estimate is the true value, the estimate is unbiased.

VARIANCE ERROR

The variance error is the first term in the expression for mean square error. Since, however, the bias error is zero, the variance error is also the mean square error in this case.

$$\therefore E \left[\hat{R}_{xy}(\tau) - R_{xy}(\tau) \right]^2 = E \left[\hat{R}_{xy}^2(\tau) \right] - R_{xy}^2(\tau)$$

Placing t on a (U, V) plane i.e.



we get

$$= \frac{1}{T^2} \int_0^T \int_0^T \left\{ E \left[x(u) y(u+t) x(v) y(v+t) \right] - R_{xy}^2(t) \right\} du dv$$

If the processes $\{x(\alpha, t)\}$ and $\{y(\alpha, t)\}$ from which the sample functions $x(t)$ and $y(t)$ are derived, are normally distributed, the expectancy bracket may be re-written thus,

$$E \left[x(u) y(u+t) x(v) y(v+t) \right] = R_{xy}(t) + R_x(v-u) \\ \times R_y(v-u) + R_{xy}(v+u+t) R_{yx}(v-u+t)$$

Hence the mean square error becomes:

$$E \left[\hat{R}_{xy}(t) - R_{xy}(t) \right]^2 = \frac{1}{T^2} \int_0^T \int_0^T \left[R_x(v-u) R_y(v-u) + R_{xy}(v-u+t) \right. \\ \left. \times R_{yx}(v-u+t) \right] du dv$$

If $\xi = v-u$ and $d\xi = du$ if the order of integration between ξ and u are reversed the above expression becomes:

$$E \left[\hat{R}_{xy}(t) - R_{xy}(t) \right]^2 = \frac{1}{T} \int_{-T}^T \left(1 - \frac{|\xi|}{T} \right) \left[R_x(\xi) R_y(\xi) \right. \\ \left. + R_{xy}(\xi+t) R_{yx}(\xi+t) \right] d\xi$$

To check that this estimate $\hat{R}_{xy}(t)$ is a consistent estimate of $R_{xy}(t)$

i.e. that the mean square error approaches zero as $T \rightarrow \infty$, integrate

the above expression from $-\infty$ to $+\infty$. This is difficult

mathematically, however, when done it proves that $\hat{R}_{xy}(\tau)$ is a consistent estimate of $R_{xy}(\tau)$.

Weakly Stationary - Autocorrelation Function

The autocorrelation function is a special case of the cross correlation function and occurs when $y(t) = x(t)$. An estimate of the true autocorrelation function is given by

$$\hat{R}_x(\tau) = \frac{1}{T} \int_0^T x(t)x(t+\tau) dt.$$

The measurement uncertainty associated with $\hat{R}_x(\tau)$ has been studied in depth by Bendat (1958), Chapter 7. Bendat has found that at $\tau = 0$ the mean square standard error is given by

$$\epsilon^2 = \epsilon^2 \hat{R}_x(0) = \frac{1}{B_N T}$$

The autocorrelation function maintains this measurement uncertainty fairly closely in the small lag region round about $\tau = 0$.

For large lag times a more complicated expression applies for ϵ^2

However, the more complicated expression gives a smaller value for the mean square standard error than does $\epsilon^2 = 1/B_N T$. Therefore it would be conservative to use the expression based upon small lag times for large lag times.

CONCLUSIONS REGARDING THE MEASUREMENT UNCERTAINTY IN CORRELATION ANALYSIS

The bias error in auto and cross correlation estimates is zero. The mean square error is therefore made up of the variance term only. Auto and cross correlation estimates are consistent.

The mean square error for cross correlation estimates is complicated at all values of τ because of the transform of the variable t into U and v . In this form it is not readily applicable to physical problems. A conservative estimate for the mean square error in cross correlation estimates would be given by the expression for autocorrelation error.

The measurement uncertainty in autocorrelation estimates may be conservatively taken as

$$\epsilon^2 = 1/B_N T$$

where B_N is the equivalent ideal noise bandwidth of the signal and T is the averaging time. This relationship is more accurate at small values of lag, (τ).

Note that the measurement uncertainty in correlation analysis is the same as that in frequency analysis. Do not confuse, however, the B 's in the expressions for ϵ^2 . It is not surprising that the expressions are similar for the frequency and time analyses. The reason being that for weakly stationary records, they are directly related by the Wiener-Khinchine equations (see Ref. 2).

MEASUREMENT RELIABILITY

Measurement reliability is a statement of confidence with which future measurements will not exceed measurements made in the past. Measurement reliability is specified for analysis data in terms of confidence limits. The confidence limit gives a measure of the range over which the true or population value will lie with a known probability. These ideas have already been introduced.

If the reliability of the analysis data has been established, the analysis data may be used to predict the magnitude of statistical properties (analysis data) that future random vibration will not exceed. Bear in mind, however, that any statement of non-occurrence will eventually be exceeded.

The actual confidence interval chosen to specify a confidence limit will depend upon the importance with which the statement of non-occurrence must be maintained. Figure will give some idea of the confidence levels that may be chosen. The example is taken from N.A.S.A. data on vibration of a Saturn Launch vehicle. Each point on the power spectrum represents the average over an ensemble of size 20.

MEASUREMENT UNCERTAINTY IN ENSEMBLE AND NON STATIONARY ANALYSIS

Mention is here made of these uncertainties, although it is not proposed to cover them in any depth.

For Ensemble Amplitudes and Time Analysis the expressions will be similar to those for single time records. Substitute in the single time record expressions for uncertainty, the expression $2/N$ for the product BT .

Ensemble frequency analysis is more difficult to implement than amplitude or time analysis. The frequency filtering operation is the difficulty that arises. One approach would be to filter the analogue time signal and then digitise and square the output. An ensemble of filtered squared digital values would arise from the collection of records. An ensemble of ensembles would be generated whose size would be the number of frequency steps needed

to cover the frequency range. The measurement uncertainty for this sort of ensemble frequency analysis has not been studied. However, an alternative method which yields similar results is to perform an ordinary power spectral density analysis for each time history record. This would generate a set of values $\{\hat{G}_{xx}(f)\}$ which could then be averaged. The difference between the two types of analysis is that the first analysis makes no assumptions about stationarity and could yield a power spectrum at different times (i.e. a time varying or non stationary frequency analysis). The second analysis involves the stationary assumption.

Therefore, for the second case the measurement uncertainty is given by $\epsilon^2 = 1/B\tau$ for each estimate $\hat{G}_{xx}(f)$. For the collection of estimates $\{\hat{G}_{xx}(f)\}$ (in accordance with statement made earlier in connection with statistical estimation theory) if $\hat{G}_{xx}(f)$ is an unbiased estimate of $G_{xx}(f)$, the true value, the average $\frac{1}{N} \sum \hat{G}_{xx}(f)$ should give $G_{xx}(f)$ without error, if the size of the ensemble is large, i.e.

The measurement uncertainty in non-stationary analysis is not yet fully understood. For the time varying non-stationary analysis the error expressions for stationary time signals should provide a reasonable approximation in most cases. As with stationary analysis, non-stationary analysis will have a compound error made up of the bias error due to finite amplitude and frequency filters and a variance error due to a finite number of ensemble members. (It has been assumed here that all non-stationary analysis will be carried out on an ensemble of records.)

APPENDIX D

DERIVATION OF THE SPECTRAL RATIOS FOR THE HUMAN OPERATOR & MAN/MACHINE DESCRIBING FUNCTIONS

Y_P in the derivation below is the Human Operator Transfer Function which is referred to as the Describing Function because the Function is based upon Statistical rather than Deterministic measures.

Y_C in the equations is the Controlled Element Describing Function and the product $Y_P Y_C$ is the combined Man/Machine Open Loop Describing Function.

The symbols used are:

i = Input Command Signal

e = Error as seen by the Human Operator

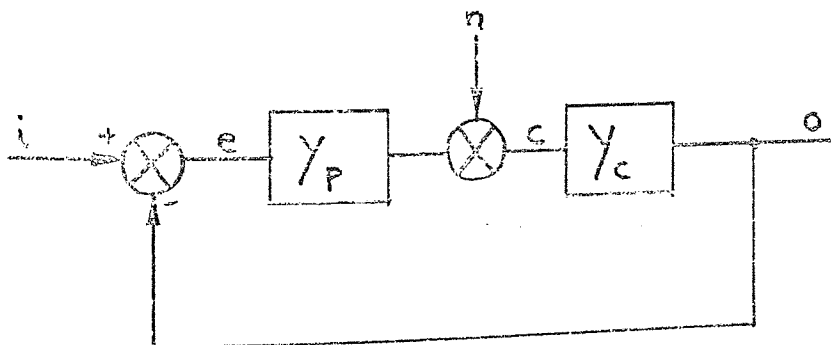
c = The Human Operators Control Movement

o = Output or State of the Controlled Element

n = Random Noise injected into the system by the Operator

The Control System Block diagram for the Man/Machine configuration

in Compensatory Tracking is shown below:



Now: $e = i - o$, $i = e + o$ & $o = i - e$. (1)

& $c = (i - o) \left\{ Y_P \right\} + n$.

also: $o = c \left\{ \frac{Y_p}{Y_c} \right\}$

$\therefore c = (i - c \left\{ \frac{Y_p}{Y_c} \right\}) \cdot \left\{ \frac{Y_p}{Y_c} \right\} + n$

$c = i \cdot \frac{Y_p}{Y_c} - c \cdot \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c} + n$

$c + c \cdot \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c} = i \cdot \frac{Y_p}{Y_c} + n$

$c(1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}) = i \cdot \frac{Y_p}{Y_c} + n$

$\therefore c = \left\{ \frac{\frac{Y_p}{Y_c}}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \right\} i + \frac{n}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \quad \text{----- (2)}$

also: $ic = \left\{ \frac{\frac{Y_p}{Y_c}}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \right\} ii + \frac{ni}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}}$

or: $G_{ic} = \left\{ \frac{\frac{Y_p}{Y_c}}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \right\} G_{ii} + G_{ni} \left\{ \frac{1}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \right\}$

If the Input and Remnant are uncorrelated then:

$G_{ni} \left\{ \frac{1}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \right\} = 0$

and: $\frac{G_{ic}}{G_{ii}} = \frac{\frac{Y_p}{Y_c}}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \quad \text{----- (3)}$

This equatio is sometimes used to define the Human Operator Transfer Function. Also:

$e = \frac{i}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} + \left\{ \frac{\frac{Y_c}{Y_p}}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \right\} n$

and $G_{ie} = G_{ii} \left\{ \frac{1}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \right\} + G_{in} \left\{ \frac{\frac{Y_c}{Y_p}}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \right\}$

Again if $G_{in} = 0$

then $\frac{G_{ie}}{G_{ii}} = \frac{1}{1 + \frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c}} \quad \text{----- (4)}$

From this Equa.

$\frac{Y_p}{Y_c} \cdot \frac{Y_p}{Y_c} = \left\{ \frac{G_{ii}}{G_{ie}} \right\} - 1 \quad \text{----- (5)}$

This is the Combined Man/Machine Describing Function.

Dividing (3) by (4) gives (ignoring the noise term);

$$Y_p = \frac{\frac{Y_p}{1 + Y_p Y_{pc}}}{\frac{1}{1 + Y_p Y_{pc}}} = \frac{G_{ic}}{G_{ie}} \quad (6)$$

This equation is considered to provide the best estimate of Y_p

From these ratio's one concludes that;

FOR CLOSED LOOP ANALYSIS; G_{ic}/G_{ii} gives the closed loop describing function $Y_p / 1 + Y_p Y_{pc}$

G_{ie}/G_{ii} gives the ratio $1 / 1 + Y_p Y_{pc}$ which if inverted and unity subtracted gives $Y_p Y_{pc}$.

G_{ic}/G_{ie} computes Y_p the Human Operator describing Function relative to the Input.

FOR OPEN LOOP ANALYSIS G_{ec}/G_{ee} gives Y_p without reference to the input signal.

Note that: In the above G_{ic}/G_{ie} cannot be easily be implemented on computer without first modifying the program to calculate the maximum value of G_{ic} for normalisation purposes. All the other ratio's which use the Auto Spectra as the denominator will normalise.

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